

Competition and Dynamic Bargaining in the Broadband Industry*

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Abstract

This paper measures the effect of horizontal mergers in the broadband industry when upstream bargaining takes time. I estimate a structural model of dynamic multilateral bargaining over interconnection fees between U.S. broadband internet providers and Netflix—the leading provider of streaming video content. Bargaining delays affect consumer welfare since Netflix streaming quality is reduced during negotiations. In my counterfactual analysis, I allow the two largest cable internet providers in the U.S. to merge before the bargaining is initiated. I find that the magnitude of the aggregate consumer welfare loss increases by 4.3% due to a longer period of degraded Netflix quality, as the merged firm exerts market power by extending negotiations. Netflix's share of the upstream surplus with the merged firm drops by 9.4%, inducing a 5.1% decrease in Netflix's probability of investing in the interconnection infrastructure that precipitated the bargaining event in the data. Prohibiting quality degradation during bargaining benefits consumers, but reduces Netflix's share of the interconnection surplus and their likelihood of investing.

Keywords: Mergers, Broadband Internet, Bargaining

JEL Classification: L41, L96, C73

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1. Introduction

Mergers between internet service providers (ISPs) are widely regarded as having the potential to reduce consumer welfare.¹ Since ISP mergers are overwhelmingly between firms serving non-overlapping markets, in practice there is no direct increase in ISP market power against consumers.² Instead, regulatory concerns focus on how increases in ISP power against the *content* side of this two-sided market spill over to consumers. A unique measurement of platform market power in this industry is the transmission or streaming quality of content to an ISP's subscribers: compared to small ISPs, a larger ISP may find it easier to degrade a content provider's streaming quality as leverage in bargaining over the prices content pays to connect to the ISP.³ If bargaining takes time to resolve, then mergers will affect consumer welfare by lengthening or shortening periods of degraded content quality during negotiations.

In this paper I evaluate the consequences of ISP mergers for consumer welfare and content provider investment, when ISP bargaining with content providers over the terms of interconnection takes time to resolve and is associated with degraded content transmission quality. I estimate a structural model of dynamic bargaining between U.S. internet service providers (ISPs) and the leading provider of streaming video content, Netflix. I then use the estimates to assess a counterfactual ISP merger and a policy intervention by the U.S. internet regulator, the Federal Communications Commission (FCC). My paper makes three contributions. First, to the best of my knowledge, my paper provides the first estimates of how ISP mergers affect consumer welfare and content provider investment. I show that a merger between Comcast and TimeWarner that was blocked in 2015 would have prolonged the Netflix quality degradation, reducing both consumer welfare and Netflix's incentive to invest in interconnection infrastructure. Second, the FCC recently ruled that ISPs could not discriminate against content by differentially reducing content providers' transmission quality; to evaluate this policy, I estimate the welfare impact of preventing streaming quality reductions during bargaining. Finally, my paper adds to a recent literature on how upstream bargaining over content affects downstream consumers (Crawford and Yurukoglu (2012), Grennan (2013)). My bargaining model recovers the upstream division of surplus and the welfare effect of bargaining from the data when upstream prices are not observed, but bargaining durations are.

The model has three types of agents: households, internet service providers, and the upstream content creator, Netflix. Each period, heterogeneous households purchase access to

¹For instance, the Federal Communications Commission (FCC) blocked the merger of Comcast and Timewarner, the two largest cable internet providers in the U.S., on the grounds that it would create a firm with enormous bargaining power vis-à-vis content providers. Becker, Carlton and Sider (2010) argues that ISP competition for subscribers is necessary to ensure positive outcomes for consumers in the absence of regulation.

²Starting in 2010 when ISP geographic footprints started to be published as part of the National Broadband Map, there have been eight large mergers. Each merger was between pairs of firms with zero geographic overlap in markets.

³Other measures of market power are also not available in this market: all ISPs provide access to all content, so there is no foreclosure, and prices in the interconnections market are missing.

the internet from the ISPs available at their residence based on ISPs' offered prices, download speeds, and whether Netflix streaming quality is currently degraded at that ISP. If households value streaming quality enough, the demand model will imply that consumers substitute between ISPs in response to degraded Netflix quality during bargaining, making it costly for ISPs to delay agreement. I assume household demand for Netflix subscriptions is growing exogenously during my sample period.⁴

The supply model explains the delays in bargaining over interconnection fees between Netflix and the ISPs. After paying an upfront development cost, Netflix can make a cost-saving, ISP-specific investment in interconnection infrastructure with each ISP, but which requires the cooperation of the ISP to implement. ISPs do not know the exact value of the investment ex ante, but instead form expectations over its distribution. An ISP can make take-it-or-leave-it offers in each period to screen Netflix's true value, but Netflix streaming quality remains degraded at that ISP until Netflix accepts an offer. Consumers' willingness to substitute away from ISPs with degraded quality will generate interactions between ISPs as they face tradeoffs between extracting investment surplus from Netflix and losing subscribers.

The objects of interest to estimate are the consumer substitution elasticities with respect to Netflix quality degradation, and the parameters governing the distribution of the investment surplus. I estimate consumer preferences for internet access and for Netflix streaming quality using individual level data on plan choices conditional on ISP choice, combined with nationwide ISP internet subscription shares and Netflix-ISP pairwise quality of service degradation. While ISPs charge higher monthly subscription fees for plans with more features, this pricing and feature schedule varies little nationwide and over time, making it difficult to estimate elasticities with market share data alone. I use fine geographic variation in the set of available ISPs and plans to identify the distribution of preferences for price, download speed, access technology, and Netflix quality of service in a mixed-logit framework with random choice sets.

To estimate the supply model parameters I use maximum likelihood. I assume that the observed data on disagreement durations is a perfect Bayesian equilibrium of the multilateral dynamic screening game. Given the consumer elasticities from the demand model and a guess of supply parameters, I numerically solve for optimal state-contingent ISP offers of interconnection fees. The optimal policies imply a joint distribution of bargaining durations for all ISPs, from which I can recover the likelihood of observing the durations that appear in the data. Since I only observe disagreement durations and not actual fees, including the profit interactions between ISPs induced by consumer quality substitution is crucial for identification of the surplus distribution. Without side-profits, two firms that face distributions of unknown surplus with the same variance but different means will adopt screening strategies with the exact same equi-

⁴This assumption is based on reduced form evidence detailed in the sections below. I relax this assumption in the online appendix.

librium distribution of agreement timings, implying mean shifters are not identified. However, with flow profits included, ISPs with large marginal losses in subscribers but long observed bargaining durations must have high means (and variances) of their surplus distributions.

From the demand curve results, I estimate that during periods of disagreement, ISP market shares decrease by about 0.5 percent on average with substantial heterogeneity across ISPs. For instance, Comcast is predicted to lose more, and higher-margin, subscribers than AT&T during bargaining, which will help match the fact that Comcast agrees more quickly in the data. The median price elasticity among all ISP-plan-quarter combinations is 2.667, substantially higher than the 0.7 estimated in Dutz, Orszag and Willig (2012). In the bargaining game, I assume a parametric distribution of ISP-specific surpluses that depends on observable features of ISPs. I estimate that there are increasing returns to scale in Netflix's cost-saving infrastructure investment: the mean and variance of the distribution of surplus are convex in the number of residences ISPs serve. Adding 10 million housing units to the median ISP's network increases the mean of the surplus distribution by approximately 3 million USD.

I examine two regulatory interventions: a merger between Comcast and TimeWarner Cable—the two largest providers of cable internet in 2013 in the U.S.—that was blocked by the FCC in 2015, and a policy that prohibits content quality degradation during bargaining.

My main counterfactual allows Comcast and TimeWarner to merge before the bargaining event begins. The model predicts that the magnitude of the aggregate consumer welfare loss due to degraded quality during bargaining increases by 4.3 percentage points after the merger. Netflix's share of the upstream surplus with the merged firm decreases by 9.4 percentage points, implying that ex ante, it would have been 5.1 percentage points less likely to make the investment in improving interconnection infrastructure that precipitated bargaining in 2013. Intuitively, the merged firm has more to gain from screening than Comcast or TimeWarner combined due to economies of scale in Netflix's interconnection surplus, but the merged firm's marginal loss in subscribers from disagreement increases only additively. The tradeoff between losing subscribers and extracting investment surplus tips towards extracting surplus, giving the merged firm the incentive to prolong negotiations. Because the marginal gains in subscribers to an ISP to concluding bargaining are larger when rival ISPs have not yet agreed, the merged firm's finer screening induces slightly more aggressive offers on average by competing ISPs, decreasing disagreement lengths and improving welfare for consumers at those ISPs.

In my second counterfactual, I prohibit the degradation of Netflix quality during bargaining. This policy is related to recent Federal Communication Commission (FCC) rules on network neutrality that ban ISPs from selectively degrading content providers' connection quality to the end consumer; my policy is slightly more general, in that it bans both content providers and ISPs from degrading the connection quality. I find consumer welfare increases, with the largest increases at large ISPs like AT&T and Comcast whose bargaining took the longest to re-

solve. However, bargaining times actually increase and Netflix's surplus share decreases, making Netflix 8.7 percentage points less likely to invest in interconnection infrastructure. This follows from the fact that in the data, marginal consumers are more willing to switch platforms (ISPs) than cancel Netflix, so prohibiting quality degradation actually removes Netflix's advantage in bargaining granted by the relative inelasticity of their subscriber base.

This paper contributes to several literatures within IO. To begin, I add to a growing literature that analyzes bargaining between firms along the supply chain. Starting with Grennan (2013), recent industry analyses have endogenized firm costs by posing a model of upstream bargaining over inputs. Papers that allow for substitution by downstream consumers in response to upstream bargaining include Crawford and Yurukoglu (2012), Gowrinsankaran, Nevo and Town (2015) and Dafny, Ho and Lee (2016). Consumer substitution is a key source of upstream bargaining incentives, and I extend these papers' insights to a dynamic setting.

A longstanding literature including Evans (2003), Evans (2010) and Evans and Schmalensee (2013) have recognized that two-sided markets—like the market for internet service—face a unique set of antitrust issues, since market power may be reflected in low prices for consumers but high prices for content providers. Empirical work including Argentesi and Filistrucchi (2007) and Chandra and Collard-Wexler (2009) focuses on the static consequences of market power in platforms; I contribute to these analyses by focusing on the dynamic consequences for content provider investment.

There is a small but growing set of tools to empirically analyze dynamic bargaining. A complete information framework for multilateral dynamic bargaining is developed in Merlo and Wilson (1995) and Merlo and Tang (2012), where delays arise if the value of surplus to be split rises over time. My model complements their analysis by allowing for delay even with constant surplus and introducing downstream competition among bargaining parties. Ambrus, Chaney and Salitsky (2016) have structurally estimate a dynamic bargaining game with incomplete information, and I extend this framework to the multilateral setting. Bargaining durations resemble joint optimal stopping problems, notable examples of which include Berry and Tamer (2006), Honoré and de Paula (2010) and Björkegren (2015). I draw on insights from this literature, especially with regards to how payoff interactions affect the set of equilibria.

Finally, the paper contributes to a literature on network neutrality and content quality reduction that has, until now, been mostly theoretical or heuristic in its analysis. Lee and Wu (2009), Becker, Carlton and Sider (2010), Economides and Hermalin (2012) Gans (2015), and Peitz and Schuett (2016) have all analyzed whether a policy of network neutrality is welfare improving, and what types of distributional impacts different neutrality policies will have. I add to this discussion the idea that apparent short-run violations of neutrality may actually be to the benefit of content providers, who may occasionally have an incentive to exert their market power through degraded quality.

The rest of the paper is organized as follows: Section 2. describes the context and motivates the bargaining delays. Section 3. details the data and its sources, and Section 4. describes reduced form patterns that inform modeling choices. Section 5. present the model of dynamic multilateral bargaining with asymmetric information, Section 6. outlines the estimation procedure and describes the obstacles to identification, and Section 7. presents the model estimates. Section 8. provides results from the two counterfactuals.

2. The Broadband Industry

2.1. Structure of the Internet

The internet is a two-sided market. On one side are consumers, who purchase access to the internet in order to consume online services and content such as email and streaming video. On the other side are the providers of services and content, such as Google and Netflix, who charge consumers either indirectly, via advertisements, or directly, via subscription fees, for using services and viewing content. In the middle are layers of firms that intermediate the relationship between consumers and content providers. In what is to follow I refer to the service/content side of the market as content providers.

Consumers and small businesses interact with "last-mile" or "edge" internet service providers like Comcast and Verizon. A consumer's choice set for wired internet service depends on which ISPs have infrastructure connected to her house, since last-mile ISPs have the exclusive right to sell service on infrastructure they own. Service is differentiated by infrastructure technology (cable, fiber optic, etc.) across providers, and by tiered menus of plans varying by monthly price and download speed in megabits per second (MB/s) within providers.⁵ By 2013, 70% of households had access to two or more wired providers offering maximum download speeds greater than 10MB/s. However, the industry is concentrated: for 91% of those consumers, at least one alternative was provided by the four largest last-mile ISPs: AT&T, Comcast, TimeWarner, and Verizon.

Netflix and other large content providers seek to connect to last-mile ISPs, and have several options to do so. The largest, like Google or Microsoft, incur a large fixed cost to install infrastructure that allows them to connect directly with last-mile ISPs at low variable cost. Others buy access from "transit" ISPs like Level3 and Cogent, who connect with last-mile providers to transmit content to consumers. Using third parties to transmit content comes with a higher variable cost, and content providers must ensure they purchase sufficient access to meet consumer de-

⁵Upload speed in MB/s, caps on how much content can be consumed in a month, and contract length are also plan features, but these are much less important: 92.5% of respondents in the 2013 Current Population Survey Internet Supplement list price, download speed, or reliability as the most important feature of service, from a list of choices that also includes upload speed, data usage caps, mobility, and bundling options.

mand. To avoid purchasing enough transit access to meet demand at peak times, content companies can also pay to upload content to so-called "content delivery networks" (CDNs)—caches of servers distributed around the country that ensure no consumer is far from a content source.

2.2. Netflix Bargaining Event

Starting in mid-2012, Netflix developed a strategy to transition from using mainly third parties to disseminate content, to using its own infrastructure. They developed a custom CDN, called Open Connect, and in so doing incurred a large fixed development and deployment cost.⁶ Open Connect would save Netflix money in two ways. First, it would allow them to save on the variable cost of using third party CDNs. As the largest online video distributor, Netflix not only paid transit ISPs for connections and the CDNs for servers, but also pursued a policy of paying the fees that last-mile ISPs charged CDNs and transit ISPs carrying Netflix content.⁷ Second, by locating the servers inside last-mile ISPs' own networks, Netflix would no longer need to ensure that it paid for sufficient bandwidth from transit ISPs to accommodate demand at peak times. With Open Connect servers located in, for instance, Comcast's network, Netflix could update the servers slowly and during off-peak times when Comcast consumers were not streaming, and therefore save on transit costs.⁸ Open Connect would allow Netflix to deliver service reliably and at lower cost.

By mid-2013, Netflix had not installed Open Connect in the vast majority of last-mile ISP networks, and had begun to report degraded quality of service to a number of U.S. ISPs. I emphasize the quality degradation for the largest two U.S. ISPs by subscriber count, AT&T and Comcast, who in 2013 collectively accounted for 43% of all U.S. broadband subscribers, in Figure 1. Starting in mid-2013, the average transmission rate of Netflix data to subscribers at these ISPs dips far below trend, and is restored after varying amounts of time. ISPs including TimeWarner (13% of subscribers) and Verizon (10.5%) also experience degradation, while Cox (5.5%) and Cablevision (3.3%) do not.

I argue that these slowdowns and their resolutions correspond to periods of bargaining disagreement over the negotiated fees for installation of Open Connect. In Figure 1 Comcast service quality is fully restored during the first quarter of 2014, which corresponds to Netflix Federal Communications Commission (FCC) filings indicating that by January, 2014, Netflix and Comcast had reached a deal on interconnection fees.⁹ AT&T service quality is only restored later:

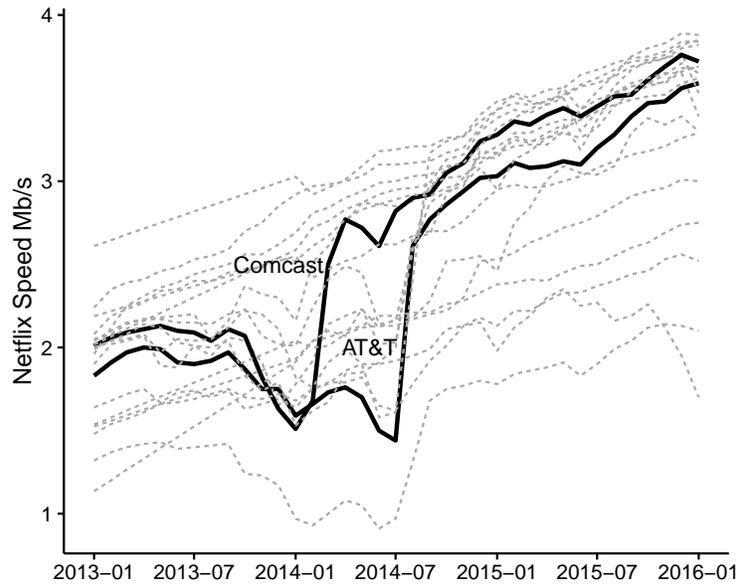
⁶Netflix Petition to Deny, pg. 49, paragraph 1. Fixed investment in R&D and deployment on the order of \$100 000 000.

⁷Paragraph 12, Statement of Ken Florance, Vice President of Content Delivery at Netflix since 2012.

⁸Netflix petition to deny, pg. 49, paragraph 2. "Open Connect...uses a 'proactive caching' method to conduct daily content updates during periods when networks are least used, such as early in the morning, to avoid congesting the network."

⁹Petition to Deny, pg. 57, paragraph 2 — pg. 58 paragraph 2.

Figure 1: Average Netflix throughput to 21 U.S. ISPs



in Netflix’s April 2014 Q-10 filing, they state that AT&T still has not agreed to Open Connect interconnection,¹⁰ but data from the Center for Applied Internet Data Analysis (CAIDA) indicates that AT&T began interconnecting with Netflix in August 2014—around the time AT&T service quality is restored. When describing the event in FCC filings at the end of 2014, Netflix notes that “none of the U.S.’s four major ISPs [had] agreed to partner with Open Connect without payment”, implying that the parties were indeed negotiating over explicit transfers from Netflix to the ISPs.¹¹

3. Data

3.1. Demand Data

Demand data is collected from several sources. An overview of the time trends between in market shares, choice sets, and plan characteristics between 2010 and 2014 is presented in Figure 2.

Market shares: Data on market shares are gathered from ISPs’ quarterly and yearly earnings reports (10-Q and 10-K) which are available for all publicly traded companies in the U.S. Total internet subscriber numbers are given every quarter; combined with auxiliary data on market sizes detailed below, these numbers imply nationwide market shares. The reports also contain ancillary data on mergers, which provide a source of variation in available plans. Some ISPs are privately held—e.g. Cox and RCN—in which case I use estimates of the subscription base from

¹⁰Netflix 2014Q1 letter to investors, pg. 5 paragraph 3.

¹¹Petition to deny, pg. 49, paragraph 2.

Leichtman Research Group. Market share movements are dominated by trends and mergers.

Plan characteristics: The menu of prices and download speeds each ISP offers are gathered primarily from the FCC Urban Rate Survey and Open Connectivity Database.¹² Where prices are missing, I collect them by hand from stored ISP frontpages on the Internet Archive Project. When the Internet Archive is unable to recover the prices—for instance, due to prices being hidden behind a localization layer—I comb ISP-specific consumer reviews on DSLreports.com. ISPs add or drop plans from their menu across different regions, but conditional on offering a plan it is advertised at the same price everywhere during the sample period. The price per megabit averaged across offered plans drops by 58% over the sample period.

Choice sets: Most consumers only have access to only one or two wired internet options. Data on what choices are available to consumers comes from the National Broadband Map (NBBM), a government initiative with data available from 2010 through 2014 which collects information at half-yearly intervals on ISP connections at the census block level. For each census block, ISPs report whether they provide service to that block and their maximum advertised speed. The maximum advertised speed truncates ISP plan offerings in that block, generating geographic variation in menus. Combined with census data on exact counts of households within each block, this data gives the weight of households across choice sets for any level of geographic aggregation. I assume that all consumers have access to satellite internet as part of their choice set. The share of consumers with access to two or more high speed (≥ 25 Mb/s) providers increases from below 20% to almost 70% during the sample.

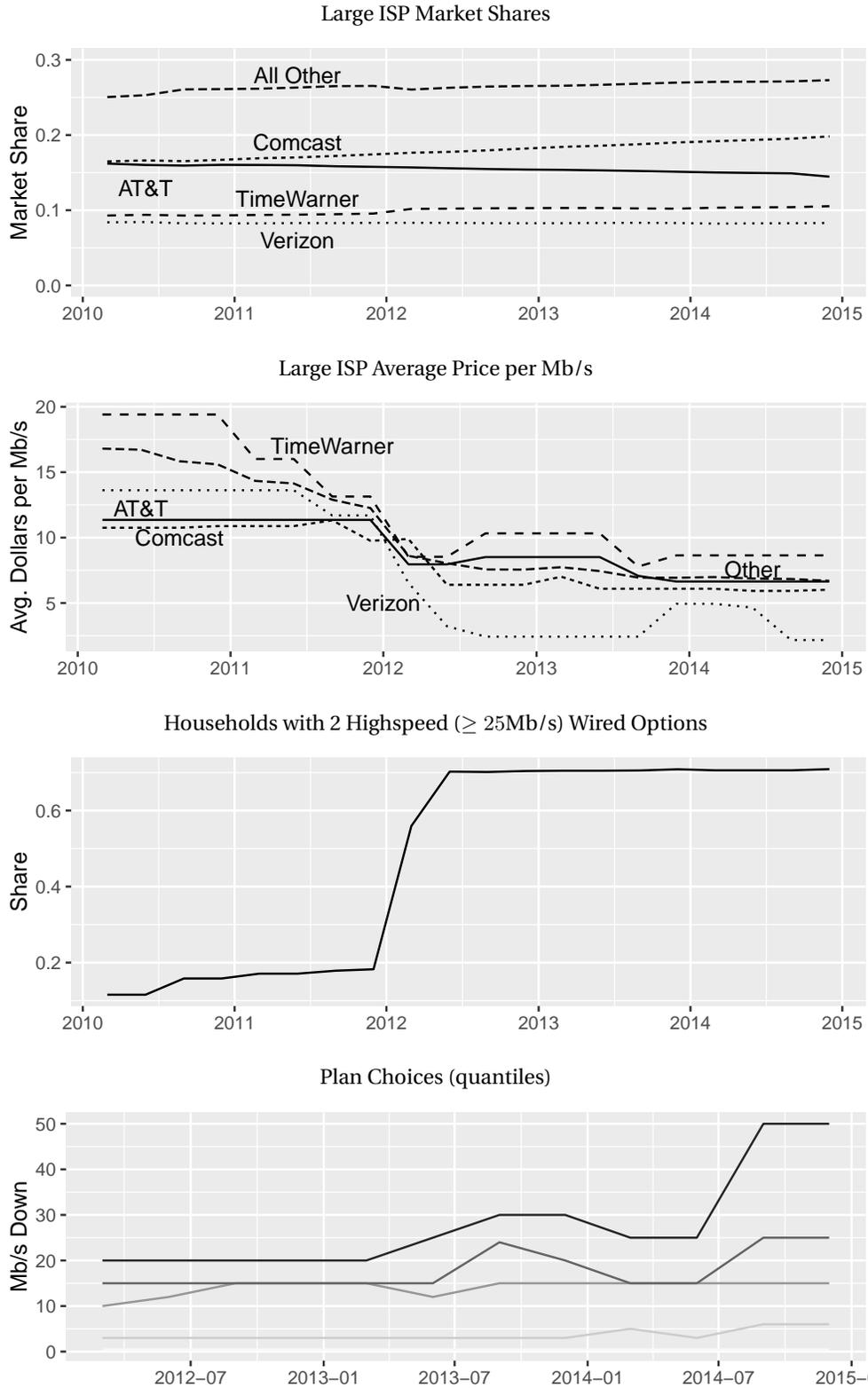
Plan microdata: I construct a time series of within-ISP plan shares using data from the FCC's Measuring Broadband America (MBBA) program. The program consists of high frequency testing data for an unbalanced panel of roughly 10 000 households from 2012 through 2014. I observe a household's ISP and their tested download speeds, which I use to back out which plan within an ISP's menu each household subscribes to in each quarter. The distribution of customers across plans changes over time, but substantial upgrading only occurs later in 2014.

Demographic microdata: The 2013 and 2014 waves of the American Community Survey (ACS) link household choices of access technology to their demographics within over 2000 Public Use Microdata Areas (PUMAs). The 2011, 2013 and 2015 waves Current Population Survey (CPS) link demographics to a variety of questions about internet usage, including use of streaming video and ISP switching behaviour.

Netflix data: I recover Netflix's quarterly subscribers, as well as the share of paid subscribers, from their Q-10 and K-10 filings. Netflix prices do not change over this time period. Whether and to what degree Netflix streaming quality is degraded is linked to bargaining disagreement durations, which I describe in Section 3.2. below.

¹²No relationship with Netflix's Open Connect CDN.

Figure 2: Time Variation in Internet Data

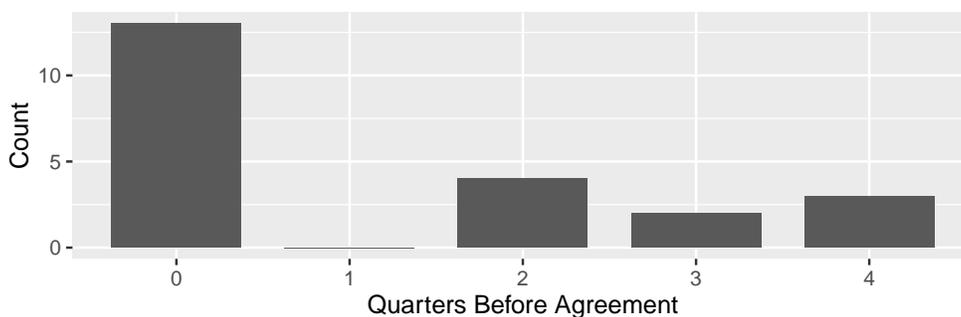


The quantile graph shows the 20,40,60 and 80% quantiles for a balanced subpanel of 1913 MBBA participants. The unbalanced panel has substantial attrition, and large additions that change the distribution.

3.2. Supply Data

Bargaining Delays: I gather data on the Netflix quality degradation event from several sources, including the Netflix data in Figure 1, data from an independent measurement company MLab, and CAIDA. In addition I draw extensively from business filings; the full data construction description, as well as the list of ISPs and associated bargaining times, is provided in Appendix C. Bargaining begins simultaneously for all U.S. ISPs in the third quarter of 2013, and lasts for between 0 and 4 quarters. I present the histogram of disagreement lengths in Figure 3. Disagreement between an ISP and Netflix affects that ISP’s customers nationwide.

Figure 3: Histogram of Disagreement Durations



4. Patterns of Consumer Substitution and Agreement Timing

4.1. Consumers switch ISPs

In this section I show that there exist marginal consumers who care about Netflix quality enough to leave ISPs that do not reach an agreement quickly.

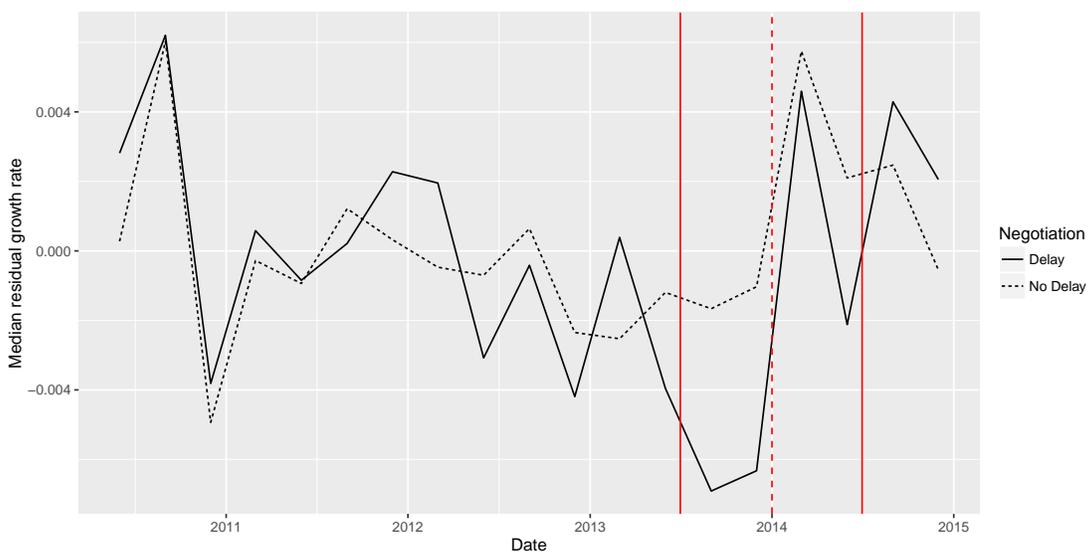
Figure 2 shows that by 2013, a majority of consumers had access to two high speed wireless options, giving them the ability to switch. The 2013 CPS Internet Use Supplement reports that 18.8% of households report having switched ISPs in the past 3 years. There is suggestive evidence that Netflix subscribers are more prone to switching ISPs on average: households with a TV-based internet streaming device were 36% more likely to switch than households without such a device.¹³ Moreover, this subset of consumers were 19% more likely to report that reliability was their reason for switching.

I use movements in aggregate subscriber shares to directly evaluate whether the slowdown affected consumer utility. Figure 4 presents median residual ISP growth rates for two groups

¹³This may be due to a difference in switching costs: streaming video users were 4.7% less likely to bundle internet with cable. It may also be an extreme manifestation of a prevalent preference for reliability: 36.7% of respondents list reliability as most important feature of a plan, greater than price (30.5%) or download speed (25.3%)

of ISPs: those that experienced a delay in negotiation with Netflix, and those that did not. The first vertical line indicates the beginning of bargaining, the dashed vertical line indicates the conclusion of the first wave of bargaining, and the final line indicates when the last ISPs agreed. A gap opens during this time period, and the median ISP that did not experience disagreement and a quality degradation grows 0.5% faster than ISPs that did. To construct these series I purge ISP specific time trends, time dummies, and control for mergers and price movements.

Figure 4: Residual ISP subscriber growth rates by whether negotiation occurred



I next run the following regression using the ISPs' reported subscriber data, fixed effects, the price per megabit of the ISP's entry tier plan, and a dummy variable indicating whether an ISP is currently in disagreement with Netflix:¹⁴

$$\Delta \log subscribers_{it} = \beta_1 disagree_{it} + \beta_2 \Delta \log p_{it} + \gamma_i + \gamma_t + \epsilon_{it}$$

If the quality of service degradation is having a persistent negative effect on the growth rate of subscribers, one would expect β_1 to be negative and significant. The coefficient on $\Delta \log p_{it}$ should also be negative, although potential endogeneity with the supply curve or unobserved demand shocks may lead it to be positive.

Results are reported in Table 1. The first four columns are simple OLS regressions with $(\gamma_i, \gamma_t) = (\gamma, 0)$. Column (1) includes all data, including time periods when a merger happens. Column (2) removes mergers; as these are large, noisy events, removing them increases significance even as it reduces the point estimates. Columns (3) and (4) repeat these exercises, but for an alternative formulation of disagreement times, which increases the duration for some ISPs

¹⁴Results are robust to different price controls.

from 2 to 3 quarters.

The fixed effect regressions (5) and (6) add in ISP and time specific dummies to control for firm level trends in subscribers as well as country-wide shocks, and (7) and (8) also control for ISPs' median price. All coefficients are negative, although estimates using the second measure for disagreement are marginally insignificant. Price has the expected negative coefficient, although it is not significant, potentially due to endogeneity with unobserved demand shocks.

The regressions confirm that firms whose Netflix quality of service (QoS) was degraded experienced a roughly 0.5% decrease in subscriber growth compared to firms that reached agreement; the results suggest an economically significant role for subscriber substitution between ISPs in response to QoS problems with Netflix. That is, I find that marginal consumers do substitute away from ISPs that are in a state of disagreement with Netflix. This suggests that not only do these consumers value Netflix streaming quality, they value it enough (and are informed enough) to leave ISPs (or at least, not sign up with ISPs) whose streaming quality is degraded.

Table 1: Consumer ISP substitution

	OLS				FE			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Disagree	-0.007*	-0.06***			-0.004**		-0.004**	
	(0.002)	(0.002)			(0.002)		(0.002)	
Disagree2			-0.008	-0.004**		-0.009		-0.009
			(0.011)	(0.002)		(0.006)		(0.006)
$\Delta \log(\text{price})$							-0.002	-0.002
							(0.004)	(0.004)
Constant	0.013***	0.006***	0.013***	0.006***				
	(0.003)	(0.001)	(0.003)	(0.001)				
Observations	541	526	541	526	526	526	526	526
R ²	0.006	0.019	0.005	0.012	0.641	0.638	0.641	0.638

The dependent variable is the log difference in ISP shares between t and $t - 1$. ***, ** and * : 0.1%, 1% and 5% significance. Standard errors are clustered at the ISP level.

4.2. Netflix Demand is Inelastic

I argue that Netflix demand during 2013-2014 is inelastic. Since I lack microdata on switching behaviour of individual Netflix subscribers and Netflix does not provide a measure of churn, I analyze the residual growth rates in Netflix aggregates from Q-10 and K-10 filings.

Figure 5 gives the residual growth rate in Netflix streaming subscribers after controlling for a time trend and seasonality. The series begins in 2012, since prior to this date Netflix aggregated its DVD rental and streaming customers. Even with controls for quarter, the growth rate exhibits periodicity; however, the pattern of subscription growth around 2014 looks very similar to that in 2013.

One possible explanation for why subscriber growth does not drop is that Netflix made more

free trial memberships available during the slowdown. That is, while Netflix does not change the sticker price of its service during this time, they may be reacting to the (endogenous) negative demand shock by offering more free trials. Figure 6 gives the residual growth rate in the fraction of streaming customers that pay for service. In 2014 the growth rate in paid subscribers is lowest during the bargaining event, unlike in 2013 and 2015 when the low point happens midway through the year. The series is very noisy, and only weakly suggests that relatively more of the growth in Netflix's subscriber base were free trial offers during the slowdown.

The evidence for substitution away from Netflix is weak. This agrees with the results in the previous section in that it points to a high consumer valuation for Netflix. Speculatively, the inelastic demand may come from the lack of similarly priced alternatives for on-demand TV and movies,¹⁵ as well as strong consumer sentiment that ISPs such as Comcast are solely responsible for network slowdowns.¹⁶ In my base model specification, I will assume that demand for Netflix is completely inelastic, but that slowdowns affect consumer valuations for ISPs. I relax this condition in the online appendix.

Figure 5: Netflix Residual Subscriber Growth

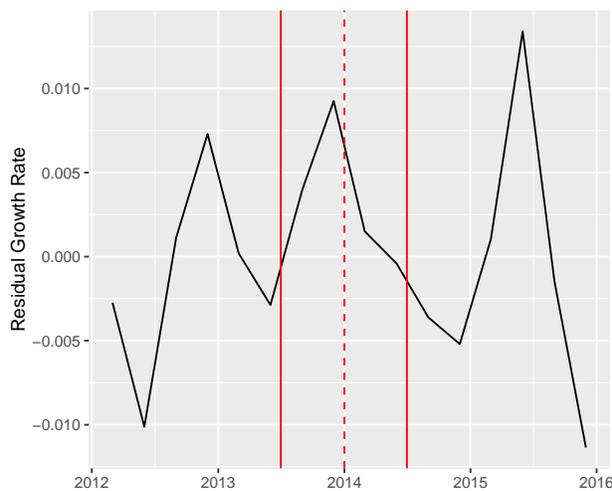
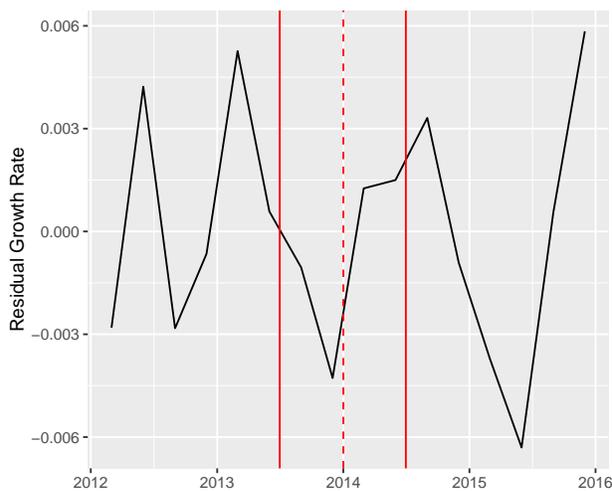


Figure 6: Netflix Res. Growth in Paid Fraction



4.3. Agreement timings correlated

If marginal consumers substitute ISPs in response to reductions in Netflix streaming quality, then agreement timings should be affected by ISP competition for subscribers. I explore whether ISPs that share more markets are more likely to have simultaneous agreement timings. Sec-

¹⁵By the beginning of 2013, DISH—the purchasers of Blockbuster—had shut down 1100 of 1500 stores, and shuttered 1450 of 1500 by 2015. A monthly Netflix subscription granting unlimited streaming was \$7.99 per month in 2013, while pay-per-view movies were anywhere from \$2.99 to \$5.99 for a one week rental. F

¹⁶Consumer Reports

tion 4.1. suggests that consumers leave ISPs that are affected by the quality of service degradation. Do ISPs recognize this substitution, and respond to it in a strategic way? That is, is the probability that two ISPs conclude agreement at the same time increasing if they compete in more markets?

To analyze this question in a reduced form way, I estimate the following specification using data from June 2013 to June 2014 inclusive:

$$agree_{it} = \beta_0 + \beta_1 comp_agree_{it} + \gamma X_{it} + \epsilon_{it}.$$

$agree_{it}$ takes a value of 1 if ISP i has concluded negotiations with Netflix at or before t . $comp_agree_{it}$ takes a value of 1 if the ISP with the greater share of overlap in i 's markets has concluded with Netflix at or before t . X_{it} is a (limited) selection of covariates.

I restrict the number of additional covariates in each specification due to limited data. I use the average offered speed of the ISP, an estimate of the fraction of the ISP's consumers that use streaming video, and an estimate of the fraction of the ISP's consumers that have bundled TV service from the 2013 wave of the CPS. Results are presented in Table 2

The $comp_agree_{it}$ variable is positive and significant in 2 out of the 3 specifications. This provides suggestive evidence that interactions between ISPs matter for agreement timings. However, this reduced form exercise cannot distinguish between whether Netflix's unobserved marginal values of agreement are more positively correlated among ISPs that compete in more markets, or whether ISPs' own strategic interactions motivate the simultaneity. In my structural model I will separate both explanations.

5. Model

There are three types of agents in the model: consumers, internet service providers, and Netflix. I assume that time is discrete, with each time period t representing three months (one quarter).

Consumers demand internet access and purchase it from the ISPs, according to their heterogeneous valuations of plan characteristics. This feature of the model predicts ISP plan shares in each period as a function of parameters, whether Netflix quality has been restored, and data on household and plan characteristics. I do not model households' choice of whether to purchase Netflix or not, but allow disutility for reductions in ISP-specific Netflix throughput.

ISPs earn profits from selling subscriptions, and may lose subscriber profits from a quality reduction in Netflix. Using estimates of the demand curve and assumptions on ISP competition and price setting, the model will predict profit as a function of the complete vector of which ISPs have restored Netflix quality of service.

On the bargaining side, Netflix and ISPs bargain multilaterally. The model predicts agree-

Table 2: Synchronicity in Agreement Timings Among Competitors

	OLS		
	(1)	(2)	(3)
<i>comp_agree_{it}</i>	0.015*	0.008	0.010*
	(0.008)	(0.007)	(0.006)
<i>speed_{it}</i>	0.16		
	(0.14)		
<i>streaming_video_{it}</i>		-0.01	
		(0.03)	
<i>bundle_{it}</i>			0.04*
			(0.01)
Observations	115	115	115
R^2	0.01	0.015	0.018

The dependent variable *agree_{it}* is a dummy that equals one if ISP *i* agrees at time *t* or before. ***, ** and *: 0.1%, 1% and 5% significance. Standard errors clustered at the unit level.

ment time and fee offer probabilities as a function of supply parameters and the estimated profit elasticities. That is, the model endogenizes agreement timings, and allows for strategic interaction between ISPs in their offers as they trade off higher fee offers against potentially losing subscribers to a rival.

I begin by describing the dynamic bargaining framework, taking the ISP subscriber profit function as given. I then describe the model of price setting and consumer demand that will serve as an input to the bargaining model.

5.1. Upstream Bargaining

Upstream bargaining is a dynamic game played between all downstream internet service providers indexed by $f = 1, \dots, F$, and the upstream content provider Netflix, indexed by N .

Time is discrete and runs from t_0 to a terminal period T .¹⁷ Bargaining is exogenously and simultaneously initiated with all ISPs by Netflix at time t_0 . At t_0 , N draws conditionally independent, ISP-specific types μ_f from a distribution $F(\mu_f|w_f, \theta_s)$ where w_f are observables.¹⁸

Netflix's vector of draws (μ_1, \dots, μ_F) is its private information. In this setting, these draws

¹⁷I postpone a discussion for why I assume a terminal period to the estimation and identification section.

¹⁸Note that the vector μ_f does not depend on the number of subscribers at the ISP. Therefore, if consumers substitute between ISPs in response to slowdowns, it does not change the size of the ISP-specific investment surplus to be split.

correspond to Netflix's ISP-specific marginal increase in surplus from installing their CDN servers in that ISP's network. w_f includes functions of observable ISP network characteristics such as their total footprint and technology, which are plausibly informative about Netflix's benefit. The eventual goal will be to estimate the parameters θ_s and use them as primitives in counterfactuals.

Actions and Timing

Starting from t_0 , within each period t :

1. Any ISP whose prior offers have not been accepted observes the history of past agreement timings, its own private information about the vector (μ_f) , and information on ISP demand shifters, and proposes a lump sum interconnection fee $\tau_{ft} \in \mathbb{R}$
2. Netflix accepts or rejects each f 's offer, $\mathbf{a}_t = (a_{1t}, \dots, a_{Ft})$ where $a_{ft} \in \{0, 1\}$. If N accepts f 's offer, N pays f τ_{ft} and realizes surplus μ_f .
3. ISPs observe the vector of acceptance/rejections \mathbf{a}_t and compete for consumers. ISPs receive flow payoffs and ISPs whose offer is accepted exit the bargaining game.

At the beginning of each period, all remaining ISPs make offers simultaneously. That is, ISPs experience both incomplete information vis-à-vis Netflix's marginal valuation, and imperfect information regarding each others' actions. Moreover, I assume that in addition to the draws (μ_f) which neither the econometrician nor f observes, there is a vector of demand shocks (ξ_{ft}) which the firms observe but which the econometrician and Netflix do not.

Netflix's problem

After exogeneously initiating bargaining at t_0 , Netflix chooses its strategy of acceptances and rejections in each period to maximize its profits. Netflix period profits are

$$\pi_{Nt}(\mathbf{a}_t, \boldsymbol{\tau}_t) = \tilde{\pi}_{Nt} + \sum_f (\mu_f - \tau_{ft}) a_{ft}.$$

$\tilde{\pi}_{Nt}$ is profit from subscriptions. In my base model I assume Netflix demand is perfectly inelastic during this time period, so $\tilde{\pi}_{Nt}$ does not depend on the vector of disagreements.

If Netflix is dynamic, their problem is to choose a vector of acceptances in each period as a function of ISP offers, the observable history of acceptances and rejections, and whatever they can infer about ISPs' information sets given histories and ISP strategies. For a history

$h_t = \{\mathbf{a}_{t_0}, \dots, \mathbf{a}_{t-1}\}$, Netflix's continuation value is:

$$V(h_t, \tau_t) = \max_{\mathbf{a}_t, \dots, \mathbf{a}_T} E_{\xi_t, \dots} \left[\sum_{t' \geq t} \beta^{t'-t} \pi_{Nt}(\mathbf{a}_t) \middle| h_t, \tau_t \right]$$

Netflix's problem is difficult for informational and computational reasons. To know whether they should accept f 's offer in a given period, Netflix must be able to forecast the next period's offer, which requires understanding how a rejection will affect the evolution of f 's beliefs about the distribution of μ_f . This issue is made complicated by the demand shock ξ , since the unobserved shock will affect f 's learning in each period in a way that is difficult for Netflix to infer. Computationally, Netflix must optimally choose a vector of actions \mathbf{a} from an action space of size 2^{23} each period (although this space shrinks as firms exit.)

To manage this complexity, in my base model I assume that Netflix is myopic with $\beta = 0$. Combined with the assumption on inelastic demand, this implies that N 's optimal strategy is separable in offers:

$$a_{ft}(\mathbf{b}, \tau) = a_{ft}(\tau_f) = \begin{cases} 1 & \text{if } \mu_f \geq \tau_f \\ 0 & \text{else.} \end{cases}$$

Before receiving offers, a myopic N would like to adopt a more sophisticated strategy, but after receiving offers they can do no better than the above.

Assuming Netflix is myopic and that demand is inelastic simplifies Netflix's strategy greatly. It also makes it trivial to invert their strategy as a function of their unobserved type, which is important for the ISPs—who will be dynamic—to be able to learn from rejected offers. Because I assume Netflix and the ISPs are bargaining over a lump sum transfer paid in the same period, there is no sense in which I am forcing Netflix to irrationally accept high offers by making them ignore a future stream of payments to ISPs. I show how to allow for elastic Netflix demand in robustness exercises.

ISPs' problem

At the end of time t , after f has observed the vector of demand shocks ξ , made offers, and Netflix has accepted or rejected each offer, firm f flow profits are written as follows:

$$\pi_{ft}(\mathbf{a}_t, \tau_{ft}) = \tilde{\pi}_{ft}(\mathbf{a}_t) + a_{ft}\tau_{ft}$$

where $\tilde{\pi}_{ft}(\mathbf{a}_t)$ is the profit earned from f subscribers at time t , which depends on \mathbf{a}_t as marginal consumers may substitute between ISPs in response to slowdowns. $\tilde{\pi}_{ft}$ will be recovered from

optimal prices and the demand curve, estimated in the next section. The second component of profits $a_{ft}\tau_{ft}$ is equal to the lump sum transfer f stands to receive if its offer is accepted.

Each ISP f seeks to maximize expected profits by choosing a best response sequence of offers. Expectations are taken over the probability of each realization of the agreement vector in each period induced by Netflix and $-f$'s strategies. f is best responding both to other ISPs' offers, as well as Netflix's optimal strategy of accepting/rejecting offers:

$$V(h_t, \boldsymbol{\xi}) = \max_{\tau_{ft}, \dots, \tau_{fT}} E_{\mathbf{a}_t, \dots} \left[\sum_{t' \geq t} \beta^{t'-t} \pi_{ft}(\mathbf{a}_t, \tau_{ft}) \middle| h_t, \boldsymbol{\xi} \right]$$

where $\boldsymbol{\xi}$ is the complete vector of demand shock realizations for all firms.

To formulate the ISPs' problem recursively, I show that given Netflix's optimal strategies and a set of ISP strategies, each ISP's information set will be a vector of (Bayesian) beliefs about the upper bound on the distributions every ISP faces. That is, f 's information will be a vector of real numbers (b_1, \dots, b_F) such that f knows $\mu_{f'} < b_{f'}$ for every f' .

First, given a set of ISP strategies, if any ISPs remain in the bargaining game at time t it must mean that their lowest offer in prior periods was rejected; given Netflix's strategy vis-à-vis f to accept any offer less than or equal to μ_f , μ_f must therefore be lower than that lowest offer. Second, notice that given ISP strategies and histories, it is possible for each ISP to construct the full information set since ISPs have no unobserved heterogeneity with respect to each other. I restrict ISPs to strategies such that if two histories of acceptance and rejection timings lead to the same information set, the optimal strategies going forward from that information set are identical. This restriction implies equivalence between histories and the information state, so that ISP optimal strategies can be solved as functions of the information state alone.

I formulate ISP f 's problem recursively as a function of the complete information state:¹⁹

$$V_{ft}(b_f, \mathbf{b}_{-f}) = \max_{\tau} \mathbb{E}_{\mathbf{a}, \tau} \left[\pi_{ft}(\mathbf{a}, \tau) + \beta V_{ft+1}(\tau, \boldsymbol{\tau}_{-f}) \middle| \mu_f < b_f, \boldsymbol{\mu}_{-f} < \mathbf{b}_{-f} \right]$$

I assume all demand shocks ξ_{ft} are perfectly forecasted, and therefore do not include it as a state variable. Moreover, I index the value functions by f and t , so that each named ISP will have its own value function to solve. Since ISPs are highly heterogeneous in their cross sectional plan offerings, market coverage, and demand shocks, as well as in the evolution of those character-

¹⁹Either the full information state, or the complete history of agreement times is necessary. The current vector of commonly observed agreement plus a firm's own information about μ_f is not a sufficient state when there are more than two ISPs. For instance, with three firms, suppose the state is such that one firm has agreed but the other two have not. Without knowing how long ago the first firm agreed, the second and third firms cannot infer each others' current strategies, which are necessary to determine the transition probabilities to next period's agreement vector. At any given time, if the first firm agreed very early, then the rate of screening by the other two firms would have been different than if the first firm agreed only recently. Each firm must be able to accurately infer what the other knows about its distribution to predict the probability of agreement in this period.

istics over time, I index value functions to avoid introducing a high dimensional state space of characteristics.

The goal in the estimation section will be to solve each ISP's optimal policies to recover probabilities of agreement timings, then use MLE to estimate the parameters of the surplus distribution, θ_s .

Alternate Sources of Asymmetric Information

To induce delay in agreement, one can introduce (1) asymmetric information as in Fudenberg and Tirole (1983), (2) irrational optimism as in Yildiz (2004), or (3) stochastic payoffs as in Merlo and Wilson (1995). I rule out optimism to keep agents rational in this first attempt at using dynamic bargaining in an industry setting. I rule out the stochastic payoff model as it requires that payoffs might increase if there is delay, which is difficult to square with this environment where quality degradation persists, and possibly grows worse, as long as there is disagreement.

An alternative to an unknown additive supply-side surplus might be unknown consumer elasticities. If ISPs and/or Netflix do not know how the marginal subscriber will substitute in response to a slowdown and they are bargaining over subscriber surplus, a delay might arise. However, if the purpose of the bargaining delay was for market participants to learn about their own subscriber elasticities, it begs the question of why ISPs and Netflix chose to learn through a costly, multi-month long quality degradation as opposed to smaller, cheaper experiments.

5.2. ISP Plan Pricing and Profits

In this section I discuss how I recover $\tilde{\pi}_{ft}(\mathbf{a}_t)$ as a function of the demand curve. I assume that ISPs take the menus of their offerings and the period's vector of agreements \mathbf{a}_t as given.

If the price elasticities from the demand curve are not greater than one, then I assume a marginal cost of zero and calculate profits as revenues from the demand curve. If the price elasticities are greater than one, then f chooses a vector of prices for its plans to maximize profits, conditional on the best response pricing of other firms. That is, f solves

$$\tilde{\pi}_{ft}(\mathbf{a}_t) = \max_{\{p_{jft}\}} M_t \sum_{j|j \in \mathcal{J}_{ft}} s_{jft}(\mathbf{p}_t(\mathbf{a}_t), \mathbf{a}_t) (p_{jft}(\mathbf{a}_t) - mc_{jft}),$$

where j is an ISP plan (price-download speed combination), M_t is the size of the market at time t , and \mathcal{J}_{ft} is the set of plans offered by f at t . s_{jft} is the nationwide demand for plan j offered by f at time t .

Taking first order conditions and inverting the system of equations to solve for marginal costs

yields the standard equation:

$$\mathbf{mc}_t = \mathbf{p}_t + \Delta_t^{-1} s_t,$$

I assume a plan has the same price nationwide. This restriction is in keeping with the data, where plan prices do not vary across geographic markets. Moreover, since I lack share data at the market level, it is impossible to recover market specific unobservable shocks which would be necessary for taking the derivative of the market specific demand curve.

5.3. Demand

This section of the model predicts demand curves as a function of parameters, Netflix quality degradation, and other data to feed into $\tilde{\pi}_{ft}$.

Flow utility

The ISPs available to a household vary broadly by geographic market m and by time t . A consumer in market m will have an individual specific choice set, reflecting the fact that not all ISPs in that market m may have infrastructure connected to that individual's dwelling. Markets are relevant to the consumer as the same ISP operating in two markets may offer different plans.

An individual i in market m at time t chooses among internet service providers f that belong to that individual's choice set, \mathcal{F}_{mt} . Each available firm f offers a menu of vertically differentiated plans $j \in \mathcal{J}_{fmt}$ which vary by market. Conditional on choosing a firm, a consumer chooses among the available j offered by f in m at time t . The indirect utility to i from choosing firm $f \in \mathcal{F}_{imt}$ is²⁰

$$u_{ifmt} = \delta_{ft} + \lambda_{ifmt} + \epsilon_{ifmt},$$

where δ_{ft} is the mean utility each consumer derives from consuming f at time t . Mean utility depends on ISP fixed effects, time, whether the ISP has reached agreement with Netflix or not, and an unobserved firm specific shock ξ_{ft} :

$$\delta_{ft} = \bar{\gamma}(1 - a_{ft}) + \bar{\alpha}_f ISP_f + \bar{\alpha}_{ft} ISP_f \times t + \xi_{ft}, \quad (1)$$

where a_{ft} is a dummy indicating whether Netflix and ISP f have reached agreement and restored quality of service. Based on the reduced form analysis, I expect that $\bar{\gamma} < 0$. Each ISP f has

²⁰I assume that switching costs between vertically differentiated products within a provider, as well as switching costs between providers or to the outside option, are all zero. Prohibiting switching costs between ISPs is done to reduce the computational burden in the supply model, since otherwise ISPs would need to keep track of consumer states.

an unobserved dimension of heterogeneity, ξ_{ft} . As in similar papers on demand for telecommunications services, this time-varying heterogeneity may reflect quality of customer service and the quality of bundled services—for example, whether a cable internet provider adds or drops channels from its TV service.

Firms are not uniquely identified with a price or download speed, so these variables do not enter δ_{ft} . Individuals choose among f 's menu of offered plans, and the effect on indirect utility is captured by λ_{ifmt} :

$$\lambda_{ifmt} = \max_{j \in \mathcal{J}_{fmt}} \{\alpha_{ip} p_{jft} + \alpha_{iq} q_{jft}\} + \alpha_{if} + \gamma_i (1 - a_{ft}).$$

p_{jfmt} and q_{jfmt} are the price and download speed associated with each plan j in the set of plans offered by f in m at time t , \mathcal{J}_{fmt} . $(\alpha_{ip}, \alpha_{iq}, \alpha_{if}, \gamma_i)$ are individual specific coefficients on price, download speed, firm technology, and whether or not there is a slowdown, respectively.

I model consumer heterogeneity $(\alpha_{ip}, \alpha_{iq}, \alpha_{if}, \gamma_i)$ as being comprised of observed and unobserved characteristics. In particular,

$$\begin{aligned} \alpha_{ip} &= -\exp\left(-\left(\mathbf{x}'_i \boldsymbol{\alpha}_p^o + \sigma_p \nu_{ip}\right)\right) \\ \alpha_{iq} &= \mathbf{x}'_i \boldsymbol{\alpha}_q^o + \sigma_q \nu_{iq} \\ \alpha_{if} &= \mathbf{x}'_i \boldsymbol{\alpha}_g^o + \sigma_g \nu_{ig} \\ \gamma_i &= \mathbf{x}'_i \boldsymbol{\gamma}^o + \sigma_s \nu_{is} \end{aligned} \tag{2}$$

where \mathbf{x}_i denotes a vector of consumer characteristics including functions of income, household size, etc. and $\{\nu_{ip}, \nu_{iq}, \nu_{ig}, \nu_{is}\}$ denote unobserved (by the econometrician) consumer tastes. Define the complete vector of characteristics for a household as $\omega_i \equiv (\mathbf{x}_i, \boldsymbol{\nu}_i)$, and the complete vector of demand heterogeneity parameters $\theta_d^h \equiv (\boldsymbol{\alpha}, \boldsymbol{\sigma})$. The full vector of demand parameters θ_d also includes the mean effect of Netflix quality degradation and ISP fixed effects, $\theta_d \equiv (\bar{\gamma}, \bar{\alpha}, \theta_d^h)$.

I assume that the unobserved tastes are distributed independent standard normal across and within consumers. As usual, ϵ_{ift} is assumed to be distributed according to the type I extreme value distribution. I borrow the functional form for the heterogeneous price coefficient from Berry, Levinsohn and Pakes (2004).

Consumers have an outside option regardless of their choice set, which reflects either non-purchase or purchase from a dialup internet provider. Indirect utility from the outside option is

$$u_{i0mt} = \mathbf{x}'_i \boldsymbol{\alpha}_o^o + \epsilon_{i0mt},$$

where ξ_{0t} is normalized to zero to fix relative utility levels, which are otherwise unidentified.

Market shares

As usual, the conditional choice probabilities for a given consumer i in market m at time t have an analytic form due to the assumption of the type I extreme value error. Given a choice set \mathcal{F}_{mt} , i 's probability of choosing f is:

$$\Pr(f_{imt} = f | \mathcal{F}_{mt}, \omega_i) = \frac{\exp(\delta_{ft} + \lambda_{ifmt})}{\sum_{f' \in \mathcal{F}_{mt}} \exp(\delta_{f't} + \lambda_{if'tm})},$$

where δ_{ft} incorporates the mean utility from choosing f in market m at time t , and λ_{ifmt} incorporates individual and market specific deviations from the mean. I assume the outside option is contained in every choice set.

Within a market m , households have a vector of probabilities of being assigned to each choice set in the market. For intuition, a market will be a geographic area defined by the census (a Public Use Microdata Area) while choice sets are observed combinations of ISPs operating in the market. The assignment probabilities potentially depend on households' observable characteristics, the moments of the distribution of observable characteristics for individuals in each choice set, and the fraction of households each choice set covers within the market. Define i 's probability of being assigned to \mathcal{F}_{mt} as $\phi_{mt}(\mathcal{F}_{mt} | \mathbf{x}_i)$, where indexing ϕ by mt incorporates the potential for the probability to depend on the distribution of observed characteristics and household weights across choice sets within a market.

To construct market shares at the national level, I aggregate across choice sets for a given i , then across i within a market, and finally across markets using potentially time-varying market weights w_{mt} :

$$s_{ft} = \sum_m w_{mt} \int \left(\sum_{\mathcal{F}_{mt}} \Pr(f_{imt} = f | \mathcal{F}_{mt}, \omega_i) \phi_{mt}(\mathcal{F}_{mt} | \mathbf{x}_i) \right) dG(\omega_i | \theta_d). \quad (3)$$

The equation Equation (3) is the model's prediction of the distribution of ISP market shares in a given t given the distribution of consumer characteristics, plan characteristics, choice sets and a vector of parameters θ_d . Matching these predicted shares to the observed shares from quarterly filings will form the basis of my estimation strategy as in Berry, Levinsohn and Pakes (2004).

The model also generates the distribution of consumer plan purchases within each firm for given parameter values. In this model, consumer heterogeneity in the valuation of prices and download speeds $(\alpha_{ip}, \alpha_{iq})$ is the only reason why within-firm plan shares are non-degenerate. Matching conditional purchase moments from the MBBA with the model's predictions will be informative about these consumer heterogeneity parameters. Because the heterogeneity parameters govern the disutility of price, these moments are essential to identify the price elasticity

of demand. The share of individuals purchasing plan j from firm f at time t is given by:

$$s_{j|ft} = \sum_{m|f \in m} \tilde{w}_{mt} \frac{\int \left(\sum_{\mathcal{F}_{mt}} \mathbf{1} \left[j = \underset{j' \in \mathcal{J}_{f_{mt}}}{\operatorname{argmax}} \{ \alpha_{ip} p_{jft} + \alpha_{iq} q_{jft} \} \right] \Pr(f|\mathcal{F}_{mt}, \omega_i) \phi_{mt}(\mathcal{F}_{mt}|\mathbf{x}_i) dG(\omega_i) \right)}{\int \left(\sum_{\mathcal{F}_{mt}} \Pr(f|\mathcal{F}_{mt}, \omega_i) \phi_{mt}(\mathcal{F}_{mt}|\mathbf{x}_i) dG(\omega_i) \right)}, \quad (4)$$

where $\tilde{w}_{mt} = w_{mt} / \left(\sum_{m'|f \in m'} w_{m't} \right)$.

6. Estimation and Identification

6.1. Demand

I use the Berry, Levinsohn and Pakes (1995) approach to estimate demand parameters, with the addition of micro moments as in Berry, Levinsohn and Pakes (2004) to help identify heterogeneity in parameters across the population and provision for choice sets that vary across individuals as in Goeree (2008). I first recover estimates of the unobserved firm (and time) specific utility δ_{ft} as a function of demand heterogeneity parameters θ_d^h . Next, for choice of instruments $Z_{f_{mt}}$ I form a GMM objective function using micro moments only to identify the heterogeneity parameters. At the θ_d^h that minimizes the GMM criterion, I recover $\delta_{ft}(\theta_d^h)$, and use it to estimate the mean effect of Netflix quality degradation and ISP fixed effects under various assumptions on the joint distribution of the exogenous component of $\delta_{ft}(\theta_d^h)$ and $\xi_{ft}(\theta_d^h)$.

From Equation (2), the demand parameters θ_d to estimate are:

Mean shifters: $\{\bar{\gamma}, \bar{\alpha}\}$

Observed heterogeneity: $\{\alpha_p^o, \alpha_q^o, \alpha_g^o, \gamma^o\}$

Unobserved heterogeneity: $\{\sigma_p, \sigma_q, \sigma_g, \sigma_s\}$.

To estimate these parameters, I match three sets of predicted moments to their sample analogues: (1) the covariance between unobserved firm-level heterogeneity and a set of instruments that shift firm markups; (2) the covariances of the observed technology type with observed consumer characteristics; and (3) the covariance of the set of instruments with the difference between the predicted conditional plan shares and observed conditional plan shares.

The first set of moments are useful for identifying the mean consumer valuation for each ISP and for the mean response to the Netflix quality degradation. Given parameter guesses, ξ_{ft} will allow the model to exactly fit the observed market shares by shifting the mean utility for each firm at each date. Depending on the joint distribution of the ISP dummies, a_{ft} and ξ_{ft} , I can

recover the coefficients $(\bar{\gamma}, \bar{\alpha})$ by OLS, FE or IV. In my base specification, interacting these ξ_{ft} with a set of contemporaneous and lagged instruments will help identify the mean shifters even in the presence of latent switching costs (see Scherbakov (2015)).

The second set of moments match observed consumer attributes from the 2013 and 2014 waves of the ACS to each consumer's chosen technology within a geographic market. These moments will be particularly useful in identifying the observed heterogeneity parameters α . If, for instance, large households choose the cable option relatively more in a market where cable download speeds are higher compared to large households in a region with poor cable options, the estimation will attribute a positive coefficient to the interaction of download speed with household size. There is substantial variation in the set of plans and technologies available across markets, mostly due to plausibly exogenous geographic variation in ISP service areas. Cross sectional and time series variation in the sets of ISPs that are still negotiating with Netflix in 2013 and 2014 identifies heterogeneous consumer valuations to the slowdown.

The final set of moments will help identify unobserved heterogeneity in the valuation of price and download speed, (σ_p, σ_q) . I match the model's prediction for the vector of nationwide conditional plan shares for each ISP observed in the MBBA data with the observed conditional shares.²¹ Variation in the menus of available plans over time and across ISPs will provide the identifying power for these parameters.

Recovering δ

I recover the δ_{ft} that rationalize the observed nationwide ISP shares according to the standard BLP algorithm. That is, given a guess for θ_d and an initial guess of the vector of $\delta^0 \equiv (\delta_{ft}^0)$, I iterate the following equation until it converges for each t :

$$\delta_t^{(a+1)}(\theta_d^h) = \log(\mathbf{s}_t) - \log(\hat{\mathbf{s}}_t(\theta_d^h, \delta^{(a)}(\theta_d^h))) + \delta_t^{(a)}(\theta_d^h), \quad (5)$$

where \mathbf{s}_t is the vector of data on national market shares at time t and $\hat{\mathbf{s}}_t(\theta_d^h, \delta)$ is the model's prediction of shares. Denote the recovered parameters by $\hat{\delta}_{ft}$, where dependence on θ_d^h is implicit.

To predict $\hat{\mathbf{s}}_t(\theta, \delta)$ requires integrating individual conditional choice probabilities across consumers and choice sets within markets. I integrate by simulation. In each m , I sample $r = 1, \dots, R$ households (with replacement) from the empirical distribution of consumer (and characteristics x) in that market, using census provided weights as sampling weights. I draw unobserved heterogeneity ν from a multivariate standard normal.

Individuals must still be assigned choice sets to construct the integral. In my baseline case, I take $\phi_{mt}(\mathcal{F}_{mt})$ to be the empirical share of households in PUMA m who are in choice set \mathcal{F}_{mt}

²¹I drop the highest plan for each provider to have an excluded category, since otherwise the moment will be flat in the parameter value.

from the National Broadband Map. For each sampled household r , I assign them a single choice set, with probabilities of each choice set dictated by $\phi_{mt}(\mathcal{F}_{mt})$. That is, if PUMA m has 40% of households in a Comcast-Verizon choice set, I will assign sampled households r to that choice set with 40% probability regardless of household characteristics. Denote r in market m at time t 's choice set as $\hat{\mathcal{F}}_{rmt}$.

The simulated market shares for firm f at time t :

$$\hat{s}_{ft}(\theta_d^h, \delta) = \sum_m w_{mt} \frac{1}{R} \sum_r \Pr(f | \hat{\mathcal{F}}_{rmt}, \omega_r, \theta_d^h, \delta_t)$$

The simulated conditional share for plan j offered by f at time t :

$$\hat{s}_{j|ft}(\theta_d^h, \delta) = \sum_{m|f \in m} \tilde{w}_{mt} \frac{\sum_r \mathbf{1} \left[j = \underset{j' \in \mathcal{J}_{fmt}}{\operatorname{argmax}} \{ \alpha_{ip} p_{jft} + \alpha_{iq} q_{jft} \} \right] \Pr(f | \hat{\mathcal{F}}_{rmt}, \omega_r, \theta_d^h, \delta_t)}{\sum_r \Pr(f | \hat{\mathcal{F}}_{rmt}, \omega_r, \theta_d^h, \delta_t)}.$$

where $\tilde{w}_{mt} = w_{mt} / \left(\sum_{m'|f \in m'} w_{m't} \right)$.

Estimating the Heterogeneity Parameters

Having recovered δ , I estimate the heterogeneity parameters using the second and third sets of moments. The second moment requires that within a market, interacting instruments with the difference in the mean characteristics of households predicted by the model to choose technology g and the mean observed characteristics of consumers who actually choose g is minimized:

$$G_R^2(\theta_d^h, \delta) = \frac{1}{NM_t} \sum_m Z_{gmt} \left[\frac{1}{n_{gmt}} \sum_{i_{gmt}=1}^{n_{gmt}} \mathbf{x}_{i_{gmt}} - E[\mathbf{x} | g(f_{imt}) = g, \theta_d^h, \delta] \right], \quad (6)$$

where n_{gm} is the number of ACS respondents reporting they use technology g in market m , and NM_t is the number of markets at time t . $g(f_{imt})$ is a function that returns the technology of firm f in market m at time t .²²

The conditional mean term in the square brackets is the model's prediction of the mean characteristics of an individual choosing technology g . I estimate it using the simulated individuals:

$$E[\mathbf{x} | g(f_{imt}) = g, \theta_d^h, \delta] \approx \frac{\sum_r \mathbf{x}_r \Pr(g(f_{rmt}) = g | \mathbf{x}_r, \nu_r, \theta_d^h, \delta)}{\sum_r \Pr(g(f_{rmt}) = g | \theta_d^h, \delta)}$$

The third set of moments use data from the MBBA testing program. These moments are similar in spirit to the second set of moments, interacting the difference between actual and

²²A firm's technology may vary, e.g. Verizon serves markets using both DSL and fiber optic cable.

predicted conditional shares with an instrument. The moments are:

$$G_R^3(\theta_d^h, \delta) = \frac{1}{\tilde{F}_t} \sum_f \frac{1}{|\mathcal{J}_{ft}|} \sum_{j \in \tilde{\mathcal{J}}_{ft}} Z_{ft} \left[\frac{n_{j|ft}}{n_{ft}} - \hat{s}_{j|ft}(\theta_d^h, \delta) \right], \quad (7)$$

where \tilde{F}_t is the number of firms present in the MBBA data at time t (with $\tilde{F}_t \leq F_t$), $n_{j|fnt}$ is the number of individuals in the data who choose j as their plan—given that they choose f at time t —and n_{nt} is the number of individuals who choose firm f at time t . $\tilde{\mathcal{J}}_{ft}$ is the set of plans offered by f at t dropping the top plan, and $|\mathcal{J}_{ft}|$ is the total number of plans offered by f at t .

Potential instruments Z_{ft} and Z_{fnt} include a firm’s own prices, technology dummies, and functions of the menu of download speeds, as well as the equivalent quantities for a firm’s competitors. The main concern is that price is endogenous to the unobserved demand shock ξ_{ft} .

Since demand shocks are firm wide, they correspond intuitively to changes in the quality of (non-Netflix related) bundled services, system-wide disruptions in reliability, or the quality of customer service. Since the shocks are observed by f , they are endogenous to contemporaneously set inputs in the firm’s profit maximization problem. I assume that only price is set contemporaneously.

The instruments Z_{ft} for G^3 that I use are variables that affect f ’s price at time t but do not depend on ξ_{ft} . I assume that the menu of download speeds evolves exogenously, according to background technological progress that enables faster speeds over the same infrastructure. Functions of competitors’ contemporaneous menu of download speeds, as well as f ’s own speeds, are valid instruments under this assumption. I also use lagged functions of the instruments to separately identify preferences from latent switching costs, as in Scherbakov (2015).

The instruments Z_{gmt} for G^2 I use are similar to the above: I include a constant, the weighted average of ISP minimum, mean, and maximum offered download speeds across ISPs of the same technology g , as well as a weighted average over all competitors’ minimum, average and maximum download speeds within each market m .

I stack $G^2(\cdot)$ and $G^3(\cdot)$ and use two-step GMM to recover $\hat{\theta}_d^h$ as the parameter that minimizes the objective function. As Hansen (1982) shows, provided that $R \rightarrow \infty$ and the ACS sample sizes go to infinity, the estimator will be consistent.

Estimating the Mean Parameters

Given an estimate of θ_d^h that minimizes the GMM objective function, I use the recovered $\hat{\delta}_{ft}$ to estimate the mean parameters $(\bar{\gamma}, \bar{\alpha})$. From Equation (1), $\hat{\delta}_{ft}$ is comprised of ISP-specific fixed effects and time trends, as well as the mean effect of the slowdown.

In my baseline case, I estimate Equation (1) by a fixed effects regression with instruments for the slowdown dummy. My timing assumption is that firms observe ξ_{ft} before they optimize

over the offer they make to Netflix. Thus, an anticipated demand shock will change the marginal value of agreement, inducing different screening and a correlation in the distributions of ξ_{ft} and a_{ft} . I instrument for a_{ft} using a subset of the markup-shifting instruments Z_{ft} described above. If f 's opponents exhibit changes in their menu of offered speeds, this change also affects f 's marginal returns to agreement with Netflix but is plausibly exogenous to ξ_{ft} .

Falsifying an Alternative Demand Model

A concern that arises from my formulation of the demand model is that it does not directly allow the Netflix slowdown to affect consumers based on their chosen download speed. One can imagine that if the slowdown is a proportional reduction in throughput to Netflix, higher download speed consumers will suffer less since they can absorb greater absolute throughput reductions before streaming quality becomes a problem.

If this concern is valid, then it introduces another reason for ISPs to delay bargaining: ISPs may throttle Netflix traffic to force consumers to purchase faster plans to upgrade their way out of the slowdown. Since faster plans are higher margin, ISPs stand to benefit, especially if consumers experience a ratcheting effect in their speed preferences over time.

I examine whether consumers change their speed upgrading behaviour at affected ISPs during periods of disagreement. I use the balanced panel of MBBA consumers to show using individual level data that consumers do not disproportionately increase their rate of upgrading at ISPs affected by the disagreement.

I run the following LPM specifications:

$$upgrade_{it} = \beta_0 + \beta_1 slow_{it} + \beta_2 \log(d_{it-1}) + \beta_3 slow_{it} \times \log(d_{it-1}) + \beta_4 t + \alpha_i + \epsilon_{it}$$

The dummy $upgrade_{it}$ takes a value of 1 if household i upgrades their speed in period t . $slow_{it}$ is the dummy indicating whether i is at an ISP currently experiencing a slowdown, and d_{it-1} is the log of i 's chosen speed last period. β_1 should be interpreted as the conditional probability of upgrading your speed if your ISP is experiencing a slowdown. Results are reported in Table 3. The coefficient on $slow_{it}$ is negative in all specifications and marginally significant with consumer and time fixed effects, suggesting that within-ISP upgrading did not increase during the slowdown.

Table 3: Consumer Speed Upgrading

	OLS	FE
$slow_{it}$	-0.010***	-0.008*
	(0.001)	(0.003)
Controls	No	Yes
FE	—	Household, Time
Observations	68832	68832
R^2	0.03	0.054

The dependent variable is a dummy indicating whether or not a household upgrades its download speed in a given month in the MBBA data. $slow_{it}$ is a dummy for whether i 's ISP experiences a slowdown in month t or not. ***, ** and *: 0.1%, 1% and 5% significance. Standard errors clustered at the unit level.

6.2. Upstream Bargaining

The supply side parameters to estimate are θ_s , a vector that governs the distribution of μ_f , $F(\mu_f|w_f, \theta_s)$. I assume that μ_f is the product of an observable function of ISP attributes w_f , and an ISP-specific shock ζ_f that is observed by Netflix but not by the ISP or the econometrician:

$$\mu_f = \exp(\mu'w_f) \exp(\sigma_\zeta \zeta_f),$$

I assume that ζ_f is distributed normal but truncated from below, and that each draw ζ_f is independent and identically distributed, which implies a truncated lognormal distribution for μ_f .

To estimate $\theta_s \equiv (\mu, \sigma_\zeta)$, I will rely on maximum likelihood. Given a guess of the parameters, I solve the optimal state contingent policy functions for ISPs, which generates a joint probability distribution of agreement timings and offers. From the data section, I only have observed agreement/disagreement timings $\{T_f\}$ so the likelihood function will have the form:

$$\mathcal{L}(\{T\}_f; \mu, \sigma_\zeta) = P(\hat{T}_1(\mu, \sigma_\zeta) = T_1, \dots, \hat{T}_F(\mu, \sigma_\zeta) = T_F)$$

For the likelihood to be well-identified, one of two situations must be true: the model may give unique predictions for agreement times for a given set of parameter values. Alternatively, the model may have multiple equilibria for given θ_s , but as long as the data is only generated by one equilibrium and I select the correct predicted equilibrium when forming the likelihood I will recover the correct parameter values.

To construct the likelihood I must find the predicted agreement times as a function of parameters, which will follow from recovering the optimal policies. I first detail what variation in the data will identify the parameters of the distribution, then describe how I recover optimal policies and in doing so, select an equilibrium to input into the likelihood.

Identification of (μ, σ_ζ)

To identify the supply side parameters, I rely on (1) cross-sectional variation in the disagreement durations, and (2) the cross sectional correlation in covariates and durations, conditional on the marginal increase in flow profits from agreement.

Estimating σ_ζ is straightforward: $\sigma_\zeta > 0$ is the only way in which the model can generate positive disagreement durations, since w_f is observed. Variation in disagreement times across ISPs that look similar in terms of covariates and flow profits can only be explained by σ_ζ .

Recovering μ is more challenging: in the standard screening model with linear utility and no side payoffs, the rate of screening is independent of the parameter μ . Increasing μ does not affect the *relative* cost of postponing agreement, so the probability of an agreement time appearing in the data remains the same for different μ . In other words, the duration based likelihood is flat in μ . However, with side payments, if high w_f ISPs that face large dips in profits if agreement is not reached have long disagreement lengths in the data, then μ must be positive and significant. This follows because μ increases both the mean and the variance of μ_f : if increasing μ lead to a pure mean shift, then μ would still be unidentified even with side payments.

Solving for Policy Functions

To form the likelihood, I must solve for optimal policies. If firms care about the states and strategies of all other firms the model quickly becomes intractable. I adopt a cutoff rule to mitigate this curse of dimensionality. First, for f consider all $-f$ such that each $-f$ overlaps at least 15% of f 's footprint. Starting with Comcast, I determine all such $-f$ —Comcast's primary competitors—and then determine the $-f$ of those primary competitors. Following this procedure until no new firms are added will form a minimal group that cares about the actions of all other members of that group, even if some members' footprints do not overlap.

The minimal group constructed in such a way is Comcast, AT&T, TimeWarner and Verizon. I assume when solving these firms' optimal policies that they pay attention to the markets that are exclusively served by members of the group. Other ISPs react to the policies of the biggest four, but do not affect the four in turn. To keep the number of ISPs a small ISP must pay attention to manageable, I assume that small ISPs assume that large ISP with which they do not overlap behave deterministically according to their durations in the data. That is, Mediacom's state space includes the information state of AT&T and Verizon, but Mediacom assumes Com-

cast deterministically reaches agreement in the third quarter of negotiations.

The equilibrium I select gives Comcast the first mover role. As in Jia (2008) or Björkegren (2015), I start from the equilibrium where all ISPs agree immediately, then do a round robin algorithm, updating the optimal strategies of each of the four large ISPs until their policy functions converge. I then recover the policy functions of the smaller ISPs give the optimal behaviour of the big four. The full set of policy functions gives the joint distribution over agreement timings for the chosen θ_s , which I use to construct the likelihood.

Terminal Dates

Since ISP value functions depend on t and not states alone, I require a terminal condition. My condition is dictated by data limitations: I have price data until the end of 2014, and the National Broadband Map extends only through the middle of 2014. I assume that bargaining terminates in December 2014 with the minimal offer, which is always accepted.²³

I predict the evolution of profits after 2014 with a flexible time series. This is necessary since firms that exit bargaining enjoy their discounted payoff into the infinite future. Given the terminal condition, I solve policy functions by iterating backwards.

Shape of Joint Density of Agreement

Each ISP will offer a strictly decreasing sequence of fees to screen Netflix's unobserved ISP specific type. Indeed, even without flow profits it is never optimal to make an offer outside the updated conditional support of μ_f , and the fact that slowdowns always hurt ISP flow profits intensifies the incentives to reduce offered fees over time.

The question remains whether an ISP will decrease its offers more quickly or more slowly once a rival has concluded agreement. This feature of the model will be important for the counterfactual merger analysis. If a rival's agreement increases the marginal benefit (from flow profits) to own agreement (supermodularity in flow payoffs), then if the merged firm screens faster so too will the other firms. If a rival's agreement decreases the marginal benefit to own agreement (submodularity in payoffs), then faster screening by the merged firm will lead to slower agreement times on average by the other firms. Figure 11 and Figure 12 give the shape of the joint density of agreement timings for two ISPs in the case where flow profits are submodular and supermodular. *Ex ante*, analytic derivations of the flow profits imply submodular profits as in Jia (2008). Full tests of the degree of submodularity are presented in the online appendix.

²³I will evaluate different assumptions on the terminal date T in future robustness checks.

7. Empirical Results

This section presents estimation results of the demand model and dynamic bargaining framework.

7.1. Demand estimates

Table 4 summarizes the key estimated parameters for the demand model. The first column refers to the ISP characteristic, and the second the household characteristic that it is being interacted with. The mean effect of disagreement is negative and significant, implying a baseline level of disutility from disagreement as reflected in the aggregate market share movements. Unobserved consumer preferences against the slowdown are large and significant while observed heterogeneity is mostly insignificant, suggesting that latent Netflix subscribership patterns are not captured by the observable covariates.

The coefficients on price and download interactions are mostly significant, and unobserved heterogeneity is very important. Recall that the price coefficient is of the form

$$\alpha_{ip} = -\exp(-(\alpha_{p1}^o + \sigma_p \nu_{ip} + \alpha_{p2}^o Inc + \alpha_{p3}^o Inc \times q_{25,Inc})).$$

Coefficient signs are intuitive: increases in income reduce the disutility of price increases, although these effects are dampened for income earners in the first quartile. For download speed, larger households have preferences for higher speeds, but this sharply drops off for households larger than 4 members.

The dispersion in unobserved preference for the price will help recover a non-degenerate predicted distribution of plan shares. This feature of the model is important not only to match the data, but to recover meaningful price elasticities. With degenerate plan shares, the price elasticity for plans with no share will be zero or infinite. Moreover, the elasticity for plans with a positive share will be too small if price changes mostly induce switching across ISPs and not switching within an ISP's own menu.

I estimate a large set of heterogeneity parameters for the cable technology dummy. Since bundling opportunities are most prevalent with cable, allowing valuation of different technologies to vary by household characteristics helps recover credible estimates of the price coefficient. I find that larger households and households headed by men draw greater utility from having a cable subscription.

The demand curve's main purpose is as an input to the bargaining game. Consumer substitution in response to a slowdown makes lengthy bargaining costly, and will identify the parameters of the distribution of unobserved surplus in combination with disagreement durations. I present the substitution patterns for the four largest ISPs in table Table 5. The experiment is to

Table 4: Demand Model Parameter Estimates

ISP char.	HH char.	Notation	Parameter	S.E.
Disagree	Constant	$\bar{\gamma}$	-0.011*	0.004*
	HH size	α_{s1}^o	-0.040	0.029
	HH size $\times 1[\text{size} \geq 5]$	α_{s2}^o	-0.076	0.040
	Speak English	α_{s3}^o	0.294	0.152
	Male head	α_{s4}^o	-0.002	0.001
	Unobs. het	σ_s	0.544*	0.247
Price	Constant	α_{p1}^o	3.698***	1.099
	Income	α_{p2}^o	0.093**	0.034
	Income $\times q_{25,Inc}$	α_{p3}^o	-0.020	0.012
	Unobs. het	σ_p	0.686***	0.203
Download	Constant	α_{q1}^o	0.027***	0.008
	HH size	α_{q2}^o	0.762***	0.202
	HH size $\times 1[\text{size} \geq 5]$	α_{q3}^o	-0.365**	0.117
	Speak English	α_{q4}^o	0.559	0.455
	Male head	α_{q5}^o	0.489***	0.145
	Age	α_{q6}^o	0.010*	0.005
	Unobs. het	σ_q	1.157***	0.330
Cable Dummy	HH size	α_{g1}^o	0.061*	0.029
	HH size $\times 1[\text{size} \geq 5]$	α_{g2}^o	0.002	0.002
	Speak English	α_{g3}^o	0.034	0.018
	Male head	α_{g4}^o	0.107*	0.043
	Age	α_{g5}^o	0.041	0.033
	Unobs. het.	σ_g	0.239	0.162

This table presents estimates of the non-linear consumer preference parameters, as well as the linear effect of disagreement $\bar{\gamma}$. Standard errors for the non-linear parameters are computed from the numerical derivatives of the GMM moments at the estimated parameters. The linear parameter estimate standard errors come from the linear IV with fixed effects regression of $\delta(\theta_d^h)$ on firm dummies and time trends, and the disagreement dummy.

turn on the disagreement dummy for each ISP in turn, assuming that all other ISPs have finished bargaining. Since demand is static, this experiment captures the marginal consumers an ISP has to lose if it does not conclude negotiations with Netflix if all other ISPs have already agreed.

In response to the slowdown, consumers at cable internet providers (Comcast, TimeWarner) mostly switch to AT&T, while consumers at non-cable providers switch to Comcast. These patterns reflect the overwhelming market presence of Comcast and AT&T as the providers with the largest footprints. This also reflects my assumptions on the functional form of utility: the best alternative to 1.50Mb/s AT&T is not 6 Mb/s AT&T because the disutility of lower streaming quality does not decrease in the purchased download speed.

Switching rates increase mostly uniformly by download speed: individuals with higher speeds switch more in response to the slowdown. The increase is especially steep for Comcast and Verizon, while at AT&T higher speed subscribers do not switch much more than lower speed ones. Since different speeds of internet are provided at roughly the same marginal cost, high speed customers are also high margin customers. Results from this table imply that Comcast and Verizon have an especially strong incentive to conclude negotiations quickly, whereas the marginal loss to AT&T from prolonged disagreement is lower.

The demand curve must be combined with a method for deriving the price-cost margin to recover profits. I compute the price elasticity for every plan in every quarter, and find that the median elasticity across offered plans in all quarters is 2.667. However, there is large variation, with an interquartile range of over 10 and almost 30% of plan price elasticities falling below one. In theory all elasticities should be above one, although elasticities for broadband below one are consistent with recent estimates in Dutz, Orszag and Willig (2012). The substantial fraction of elasticities falling below one precludes using a demand system inversion to recover marginal costs, so I assume a marginal cost of zero and take the evolution of prices over time as exogenous in my base case for estimating the supply parameters.

To assess model fit, I regress the model predicted conditional plan shares against actual plan shares from the MBBA data. I recover an R^2 of 0.76, and across the entire sample I cannot reject the null that the slope is one. I also find that the model predicts degenerate within-ISP plan shares in only 6% of quarter-ISP combinations, and never for any of the largest four ISPs. To indirectly validate the predictions in Table 5 and verify that the model does not do poorly during the periods of the slowdowns, I calculate the quarter by quarter R^2 and slope estimates for the regression of predictions on data.²⁴ I find that the R^2 has a slight inverted u-shape, with the highest values during 2013. The model's good performance even during the slowdown is encouraging, although it also likely reflects the fact that unobserved price and download heterogeneity in the model does not vary over time but plan characteristics are slowly improving.

²⁴Unfortunately the MBBA does not track well consumers who switch ISPs, so there is no way to directly verify in the data if the substitution patterns in response to a slowdown predicted in hold.

Table 5: Consumer Substitution in Response to Streaming Quality Degradation

ISP	Plan	% Switch	Best Alt.	Best Plan	% Switch cond.	% Out cond.
AT&T	1.50	0.31	Comcast	3	22.27	46.88
AT&T	6	0.29	Comcast	3	23.20	40.30
AT&T	12	0.32	Comcast	3	22.87	40.38
AT&T	18	0.34	Comcast	50	27.08	36.15
AT&T	24	0.34	Comcast	50	26.46	22.00
AT&T	45	0.36	Comcast	50	21.95	9.28
Comcast	3	0.38	AT&T	1.50	16.27	46.90
Comcast	50	0.52	AT&T	24	20.41	25.14
Comcast	105	0.65	AT&T	45	30.35	6.43
TimeWarner	1	0.43	AT&T	1.50	20.56	49.06
TimeWarner	15	0.43	AT&T	24	21.18	37.28
TimeWarner	20	0.45	AT&T	24	40.14	10.64
TimeWarner	50	0.53	AT&T	45	20.72	16.04
Verizon	0.50	0.29	Comcast	3	18.41	46.68
Verizon	7.10	0.25	Comcast	3	11.76	31.30
Verizon	50	0.35	Comcast	50	25.62	29.29
Verizon	75	0.43	Comcast	50	21.66	10.28
Verizon	150	0.47	Comcast	105	37.68	2.52

The experiment in this table is to turn on the disagreement dummy for each ISP in turn, assuming that all other ISPs have agreed already. I use data from 2013-06-01. The Best Alt. and Best Plan columns gives the ISP and plan to which the most subscribers switch conditional on switching. The last two columns give the percent of individuals who switch to the next most popular plan, and the percentage who switch to the outside option, conditional on switching.

7.2. Upstream Bargaining Estimates

The demand estimates from the previous section are combined with the dynamic multilateral bargaining assumptions and disagreement durations to estimate the distribution of Netflix's marginal value of agreement vis-à-vis each ISP. Estimates of the coefficients of the mean shifting variables w_f , as well as the standard error of the dispersion of information σ_ζ , are presented in Table 6.

Table 6: Supply model estimates

	MLE		
	(1)	(2)	(3)
σ_ζ	1.538	0.204	0.171
Pass-by	—	7.311	6.618
Pass-by squared	—	—	0.709
DSL dummy	—	0.456	-0.130

The DSL dummy takes a value of 1 if the ISP's primary transmission technology in the last-mile is DSL. Pass-by is the number of houses connected to the internet by the ISP in hundreds of millions. Standard errors coming soon.

The most parsimonious version of the model in column (1) features σ_ζ only, so that $\mu_f = \exp(\sigma_\zeta \zeta_f)$, with ζ_f unobserved by the ISPs. With no shifters w_f , the ISPs whose marginal loss in subscribers is largest would be predicted to agree most quickly. In the data the largest ISPs have the greatest marginal loss in subscribers but also take the longest time to agree, which will require the presence of the shifters w_f to rationalize.

Since the number of ISP that bargain is limited ($N = 22$), I restrict the number of covariates to functions of pass-by—the number of houses each ISP has infrastructure connected to—and the ISP's connection technology. My preferred estimates are column (3), where I allow a quadratic term in pass-by. The quadratic term allows for increasing or decreasing returns to scale for Netflix in installing the interconnection technology; the positive estimated coefficient in specification (3) suggests increasing returns to scale, so that Netflix's cost savings from making the investment increase in the size of the ISP's network. This agrees with Netflix's own statements increasing returns to the value of interconnection with larger ISPs.²⁵

²⁵"...our actual costs rise with every additional point of interconnection...Moreover, interconnecting at more locations can be less efficient, and thus more expensive, for Netflix. With each additional location, the costs associated with placing and maintain each location—e.g., servers, routers, rent, cabling costs—are spread across a smaller amount of capacity." Paragraph 20, Statement of Ken Florance, Vice President of Content Delivery at Netflix since 2012.

All else equal, the larger a provider's network, the greater the value in interconnection: evaluated at mean pass-by, increasing the number of households passed by 10 million increases the mean of the surplus distribution by 3.2 million. The positive estimated quadratic term is especially important for the merger counterfactual, since it implies that interconnecting with Comcast and TimeWarner's networks separately is not as valuable as interconnecting with the single, merged provider whose network is equal in size to the sum of the other two.

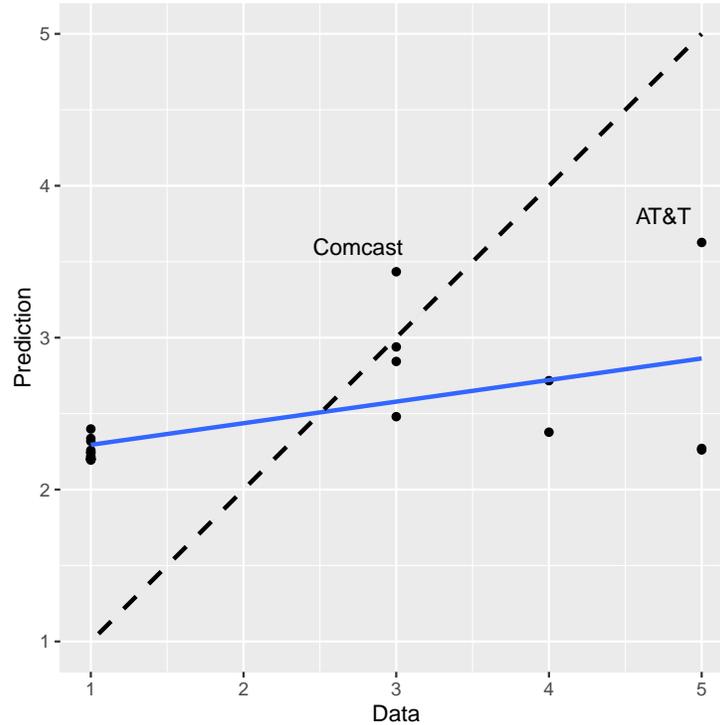
The model predicts the expected share of bilateral surplus that accrues to Netflix and each ISP. To see this, suppose that ISP f makes offer τ_{ft} that is accepted by Netflix. The model predicts that the surplus is distributed $F(\mu_f | \mu_f < \tau_{ft-1}; w_f, \theta_s)$, so that Netflix's expected surplus share in that period is $E_{\mu_f}[(\mu_f - \tau_{ft}) / \mu_f | \mu_f < \tau_{ft-1}]$. Integrating over the marginal distribution of agreement timing offers for a given ISP gives Netflix's expected share for that ISP. Netflix's *ex ante* expected share ranges between 28% against AT&T and 33% against Comcast. Part of this share is driven by modelling assumptions: myopia substantially increases Netflix's willingness to accept high fee offers, decreasing their share of the surplus. However, variation across ISPs will mostly reflect differences in ISP subscriber elasticities, which reflects real variations in Netflix's bargaining power.

Since the surplus is only divided after a delay, from the *ex ante* perspective a certain fraction of surplus will be lost by pushing its realization into the future. On average lost surplus is roughly 12%, with a high of 14% against AT&T.

As a byproduct of estimation, the model predicts the exact offers that are made in the period where Netflix agrees. These fees are potentially of interest to other large content providers if the practice of paying terminating fees to last-mile ISPs becomes more prevalent in the future. I find lump sum transfers of 19.0 million USD to AT&T, 12.64 million to Comcast, and 3.89 million and 4.90 million to TimeWarner and Verizon, respectively. Smaller ISPs recoup fees around 1 million USD, with the lowest fees going to Mediacom and Brighthouse—two small ISPs whose bargaining only resolved after a 4 quarter delay. The largest ISPs and Netflix earn billions of dollars in yearly revenue, so these interconnection fees are comparatively small.

I assess the model fit by plotting the predicted time to a successful offer against the actual time. The correlation is positive and significant and while an unweighted regression is nowhere near the 45 degree line, the model does perform better for the larger ISPs. The ISPs that agree immediately in the data are predicted to take just over a quarter longer to agree in the model. Since their pass-by is so small, their average time to agreement mostly reflects the estimate of σ_ζ . Recall that the observable component on μ_f can increase dispersion and time to agreement, but cannot ensure positive probabilities of delay on its own. Since the estimate of σ_ζ must also rationalize the presence of small ISPs who still take time to bargain, it will tend to overpredict delays for ISPs that agree quickly in the data and underpredict delays for ISPs that take more time. This is not a problem with the model however, as differences in agreement times for similar

Figure 7: Supply Model Fit



ISPs simply reflects variation in the draws of μ_f .

8. Counterfactuals

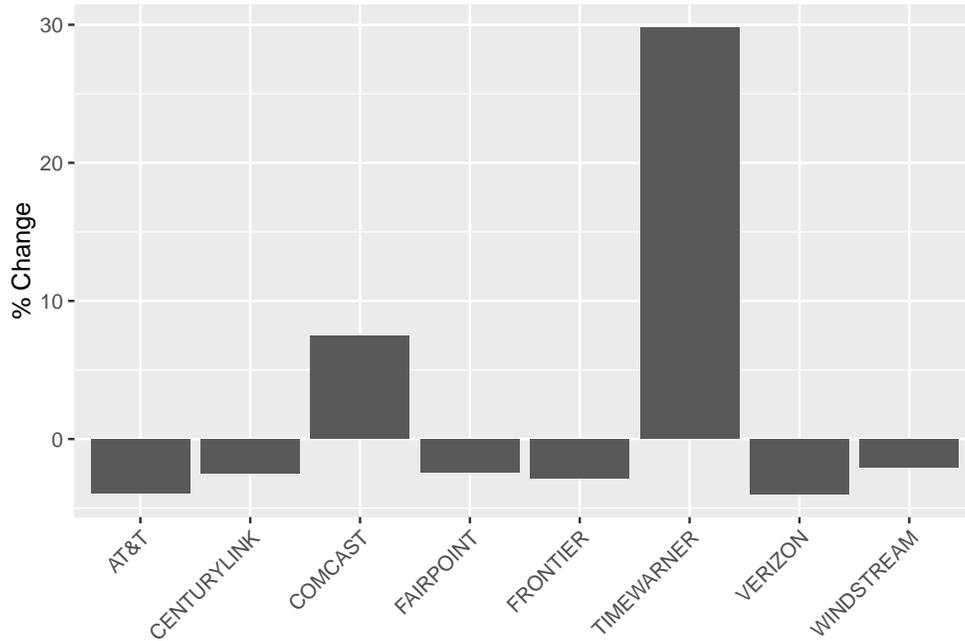
8.1. Comcast-TimeWarner merger

With both the supply and demand estimates in hand, I turn to the paper's main counterfactual—the effect of a merger between Comcast and TimeWarner.²⁶ There are three dimensions on which to evaluate the merger. Do bargaining times change? Does consumer welfare change? And do the *ex ante* predicted splits of the upstream surplus change? The first two criteria speak directly to dead weight loss and consumer welfare, which are the typical targets for judging the value of a merger. In this case the delay in agreement generates a deadweight loss, both because streaming quality degradation directly reduces welfare and because it pushes realization of surplus to the (discounted) future. The third criterion speaks to the increased degree of extractive power held by the larger merged firm which affects the split of surplus in upstream bargaining, but not total surplus.

I first examine how bargaining times vary across all ISPs. I plot the change in the *ex ante*

²⁶Since there are many price elasticities that are less than one, in my base case I assume that the merged firm simply takes on the plan characteristics and pricing of Comcast instead of doing a competitive repricing.

Figure 8: Bargaining Delay Changes



expected bargaining times in Figure 8 for ISPs that experience a non-negligible change. The merged ISP has an *ex ante* expected disagreement length of 3.69; while this is not a huge percentage increase in duration length for customers who were previously with Comcast (3.43 quarters), it is quite large for former TimeWarner customers (2.84 quarters). Other ISPs that experience changes in bargaining duration due to the merger uniformly experience slight decreases. From Table 7, notice that these ISPs are all ones that have a great deal of market overlap with Comcast and/or TimeWarner.

The quadratic term in the mean-shifting component of the surplus distribution drives the result. The merged firm has the combined footprint of Comcast and TimeWarner, but the mean and dispersion in μ_f increase more than linearly. Meanwhile, since there is 0% market overlap between these ISPs, the merged firm's marginal loss in subscriber revenue as a result of disagreement is simply the sum of Comcast and TimeWarner's marginal losses. A greater mean and dispersion in unobserved surplus combined with a relatively lower marginal cost to Netflix quality degradation implies that the merged firm will pursue a finer screening strategy in the new equilibrium. Since profits are submodular in the agreement vector, the ISPs that compete with the merged firm will screen slightly faster on average to capture the greater expected marginal gain of being the first to agree.

I estimate aggregate consumer surplus under three scenarios: without any quality degradation, with the joint distribution of quality degradation predicted by the model, and with the counterfactual joint distribution of degradation when Comcast and TimeWarner are allowed to

merge. I use the standard closed-form logit formula to derive individual surplus and then aggregate across consumers and markets.²⁷

The *ex ante* aggregate consumer surplus decreases by 0.51 percent moving from no quality degradation to the expected degradation lengths produced by the fitted model in Figure 7. Aggregate surplus decreases by 0.54 percent after the merger, a roughly 4.3 percentage point increase in magnitude. The aggregate decrease masks heterogeneity in responses across choice sets. The 20 percent of households facing choice sets that previously included TimeWarner see the magnitude of their aggregate welfare loss increase by 17.6 percentage points (to 0.60 percent), while the 34 percent of households in choice sets that include Comcast see their aggregate welfare loss increase by . Households facing choice sets that include a DSL or Fiber provider that share a large footprint with the merged firm, but that do not include the firm itself actually enjoy an increase in aggregate surplus of 0.3 percent due to slightly shorter expected slowdown durations.

The merger affects the amount of surplus that is lost due to delays. The substantial increase in *ex ante* predicted disagreement times compared to the base case implies that 16 percent of the upstream surplus is expected to be lost in the bargaining game between Netflix and the merged ISP, an increase of almost 30 percentage points from the case with no merger. Reductions in *ex ante* disagreement times for other ISPs imply that the share of surplus that is lost in bargaining decreases slightly from 12 to 11 percent on average. Netflix's *ex ante* expected share of the surplus with respect to the merged firm decreases to 29 percent, from an aggregate (combined Comcast and TimeWarner) share of 32 percent, a decrease of just over 9 percentage points.

The split and size of upstream surplus matters for Netflix's decision to invest in the CDN infrastructure. In particular, if the bargaining environment grows more adverse then there will be a hold-up problem, as either bargaining delays or greater ISP bargaining power prevents Netflix from realizing all of the surplus from its investment. I investigate the effect of the merger on Netflix's incentive to invest by analyzing how the merger changes the *ex ante* distribution of aggregate surplus net of transfers to ISPs. I will compare these distributions of aggregate surplus to an estimate of Netflix's upfront investment in the CDN.

I assume that Netflix first draws ISP specific surpluses, then pays the R&D cost.²⁸ From the model, I can recover the *ex ante* distribution of aggregate surplus net of transfers accruing to

27

$$\hat{W}_{rmt} = \frac{1}{\hat{\alpha}_r} \log\left(\sum_{f \in \mathcal{F}_{imt}} \exp(\hat{\delta}_{ft} + \hat{\lambda}_{rfmt})\right), \quad \hat{W}_t = \sum_m w_{mt} \frac{1}{R} \sum_r \hat{W}_{imt}$$

²⁸Intuitively, this reflects the idea that Netflix does a cost benefit analysis before committing to the project.

Netflix given optimal firm strategies in both the original and merger cases:

$$\sum_f \beta^{T_f} (\mu_f - \tau_{fT_f}(\mu_f)) \sim G, \quad \sum_f \beta^{T'_f} (\mu'_f - \tau'_{fT'_f}(\mu'_f)) \sim G',$$

where μ_f is random, $T_f(\mu_f)$ is the random stopping time that depends on the draw μ_f , and the variables with primes denote the same quantities drawn under the counterfactual industry structure where Comcast and TimeWarner merged.

I do not observe the initial fixed cost Netflix must expend on research and development before it can begin offering to install the CDN in ISP networks. I recover an estimate of the maximum possible fixed cost: given the actual observed agreement times, the model predicts the actual offer and an upper bound on each ISP-specific surplus as the previous period's rejected offer.²⁹ The sum of the upper bounds minus the actual offer (appropriately discounted) gives Netflix's maximum possible surplus from negotiation, which is an upper bound on its fixed cost. I estimate this quantity as 9.03 million USD.

I plot the histograms of G and G' in Figure 9. The distribution of net surpluses without the merger first order stochastically dominates the distribution of surpluses with the merger. Given the fixed cost, the leftward shift in the distribution of surpluses reduces the probability that Netflix invests in the CDN by 5.1 percentage points, from 19.4 percent to 18.4 percent. These probabilities of investment are low because the fixed cost is an upper bound.³⁰

8.2. Prohibiting Quality Degradation During Bargaining

In this section I analyze how outcomes would change if bargaining was not linked to quality degradation of the content provider's content.³¹ Insofar as degradation is a choice made by one or both bargaining parties to gain leverage in the bargaining game, prohibiting quality degradation can be implemented.

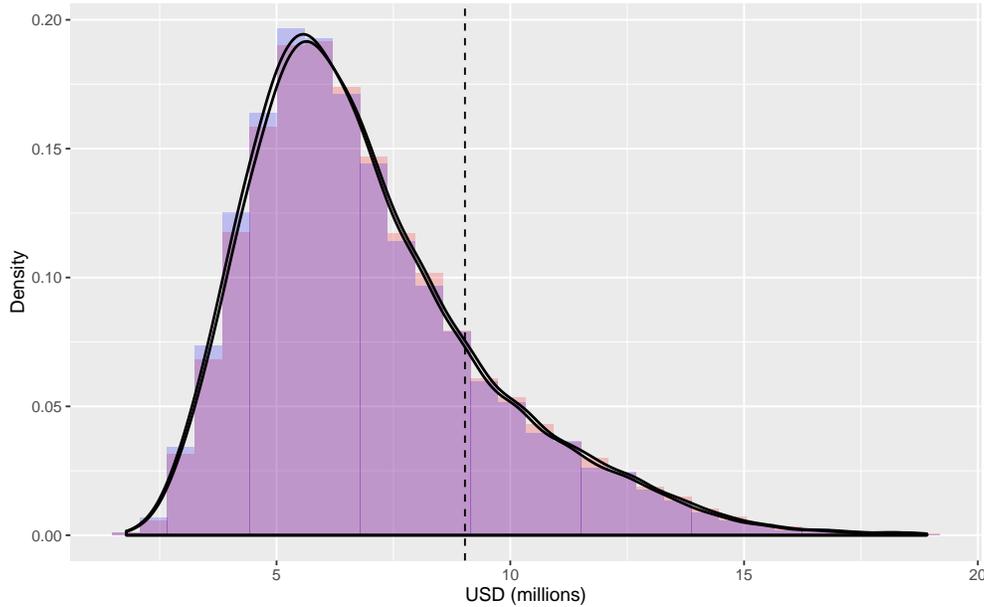
This constraint is a more general form of what is referred to as "network neutrality." Network neutrality is a policy—recently affirmed by the FCC—that ISPs cannot differentially throttle or privilege the transmission of different content, and is partly intended to prevent them from exercising their market power to extract payments from content providers. In my model, the marginal consumer will switch ISPs to prevent Netflix quality degradation, so it is Netflix that has the incentive to degrade quality to induce faster screening by ISPs. The idea that a content provider with market power may degrade quality to extract concessions in bargaining with ISPs

²⁹Since the surplus distribution is unbounded from above, for ISPs that agree in the first period of bargaining I assume that maximum surplus is the 97.5% quantile of their surplus distribution.

³⁰The lower bound on fixed costs is zero.

³¹Explicitly regulating the rules of bargaining is not an unusual counterfactual. In Gowrinsankaran, Nevo and Town (2015), the authors analyze the effect of a counterfactual hospital merger if the constituents are still forced to bargain separately.

Figure 9: Distributions of Netflix Surplus Net of Transfers



The blue histogram is the distribution of aggregate net surplus accruing to Netflix under the merged regime, the red histogram is the distribution under the base regime. Distributions are recovered by sampling the vector ζ_f $R = 10000$ times and deriving the surpluses net of transfers from ISP optimal strategies under each regime.

is a new one in the network neutrality literature, since in typical theory models ISPs are assumed to be monopolists.³²

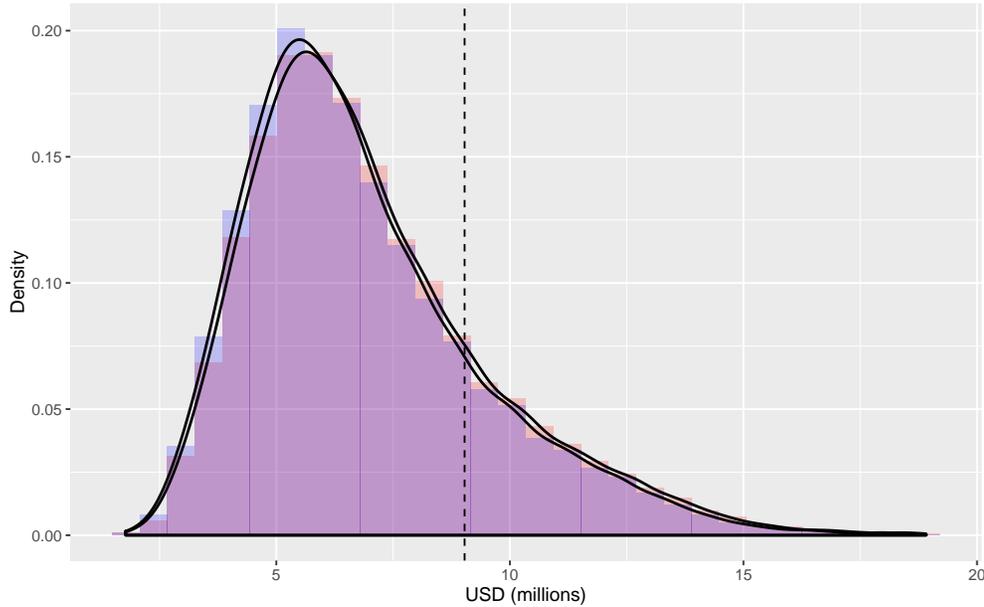
Preventing quality degradation during bargaining increases aggregate welfare by 0.54 percent (the reverse of the 0.51 percent decrease in the previous section). ISPs with longer delays such as AT&T and Comcast see consumers in their choice sets have the largest increases in welfare of 0.58 percent and 0.56 percent respectively.

Netflix's share of the *ex ante* bargaining surplus goes down and the delay to agreement increases. Without the marginal loss in subscribers affecting the screening tradeoff between finer offers and a higher payout, and coarser offers and a quicker payout, the ISPs screen much more slowly. I summarize the effects in the graph of *ex ante* distributions of consumer surplus versus the fixed investment cost Figure 10.

The probability of investment declines by 8.7 percentage points under the regime where quality degradation is prohibited during bargaining. In general, there is a two-sided holdup problem since infrastructure investment will only come on line if both sides in a bilateral monopoly agree to its installation. My setup analyzes a major investment by Netflix, and shows that the holdup problem grows more severe when quality degradation is prohibited since the marginal

³²See, for instance, Economides and Hermalin (2012)

Figure 10: Distributions of Netflix Surplus Net of Transfers



The blue histogram is the distribution of aggregate net surplus accruing to Netflix under the regime with no quality degradation during bargaining, the red histogram is the distribution under the base regime. Distributions are recovered by sampling the vector ζ_f $R = 10000$ times and deriving the surpluses net of transfers from ISP optimal strategies under each regime.

consumer is more elastic with respect to ISPs than Netflix subscriptions.

9. Conclusion

This paper combines a model of dynamic multilateral bargaining with demand for broadband internet to evaluate the effect of a proposed merger. I estimate an industry model of demand, plan choice, pricing and interconnection bargaining using data on plan prices, consumer choice sets and bargaining delays between major U.S ISPs and the leading purveyor of streaming video content, Netflix. To endogenize disagreement times as a function of ISP competition, I develop a dynamic bargaining model with asymmetric information. ISPs make take-it-or-leave-it offers to learn about Netflix's surplus from interconnection, while simultaneously competing for subscribers who value Netflix quality of service. Intuitively, if delaying agreement over interconnection degrades quality of service to subscribers, then the opportunity cost of lost subscriptions identifies the parameters of the distribution of unobserved surplus.

I structurally estimate the model, and find that a proposed merger between TimeWarner and Comcast that was challenged by the Federal Communications Commission would have slightly raised interconnection fees and bargaining length, increasing the magnitude of aggregate con-

sumer welfare loss. Consumer aggregate welfare decreases by 0.54 percent in the merger case and 0.51 percent in the base case—an increase in magnitude of 4.3 percentage points. Netflix's share of upstream surplus decreases by 9.4 percentage points, which exacerbates a holdup problem in investment in internet infrastructure. A regime where quality degradation is prohibited during bargaining improves consumer welfare by 0.51 percent compared to the base bargaining case. I find that Netflix has a lower share of surplus in this regime, and is 8.7 percentage points less likely to invest in the infrastructure project whose surplus is being bargaining over.

Future directions for this research include estimating a richer model of demand, incorporating individuals' preferences for different types of streaming content in two stage utility function as in Crawford and Yurukoglu (2012). By incorporating substitution between types of streaming content into consumers' decisions, the information encoded in the bargaining event could be used as a laboratory to test out the effects of vertical integration between ISPs and video streaming services.

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A Auxiliary Figures and Tables

Table 7: Choice set stats

Market	Fraction	Cumulative Fraction
AT&T-COMCAST	0.154	0.154
AT&T-TIMEWARNER	0.098	0.252
COMCAST-VERIZON	0.091	0.343
CENTURYLINK-COMCAST	0.067	0.410
TIMEWARNER-VERIZON	0.053	0.463
AT&T-CHARTER	0.051	0.514
CABLEVISION-VERIZON	0.044	0.558
AT&T-COX	0.030	0.588
HUGHES	0.026	0.614
CENTURYLINK-COX	0.026	0.639
RCN-VERIZON	0.019	0.658
COMCAST-FRONTIER	0.017	0.675
CENTURYLINK-CHARTER	0.016	0.691
FRONTIER-TIMEWARNER	0.016	0.707
AT&T-SUDDENLINK	0.013	0.721
AT&T-BRIGHTHOUSE	0.013	0.734
CENTURYLINK-TIMEWARNER	0.013	0.746
COX-VERIZON	0.012	0.759
AT&T-WIDEOPENWEST	0.012	0.771
BRIGHTHOUSE-VERIZON	0.011	0.782

Table 8: Conditional means of demographics from the ACS

	Technology	Income	HH Size	Speak English	Male Head	Age of Head
Uncond.	—	89.220	2.193	0.970	0.591	51.860
Internet	No	43.774	1.789	0.939	0.512	61.543
	Yes	97.801	2.489	0.979	0.614	51.098
Tech	Cable	100.704	2.460	0.979	0.606	50.446
	Dialup	63.112	2.067	0.975	0.590	61.992
	DSL	90.060	2.531	0.979	0.625	52.297
	Fiber	121.480	2.607	0.984	0.636	51.684
	Mobile	79.529	2.495	0.970	0.602	48.401
	Satellite	77.424	2.440	0.981	0.631	54.976

Figure 11: Submodular profit interactions

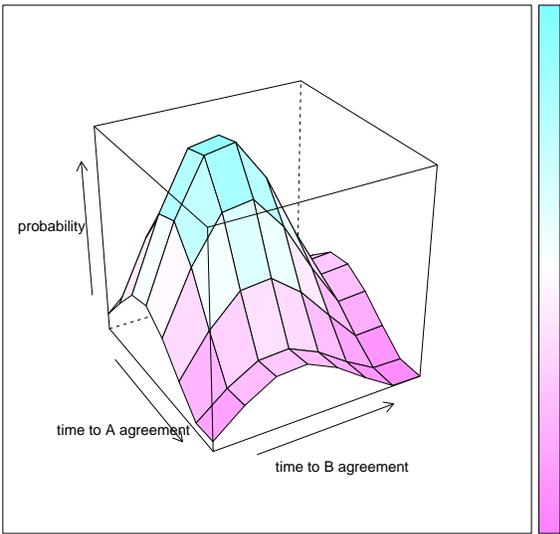
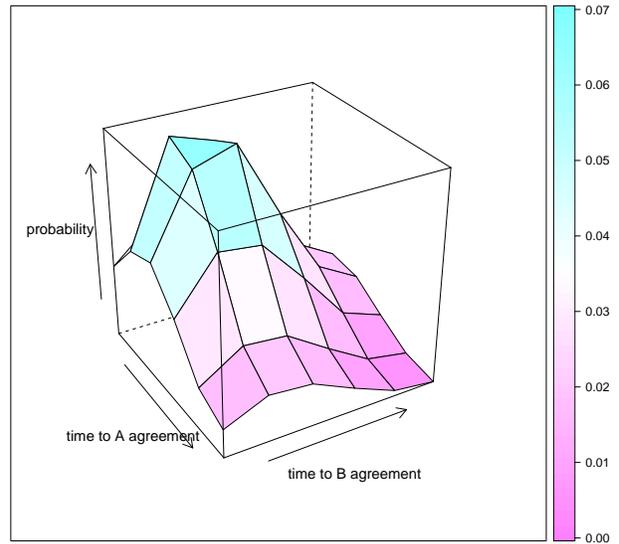


Figure 12: Supermodular profit interactions



B Network neutrality

Early architects of the internet promoted a design such that all content could be equally accessible by all consumers, regardless of where the content originates or which last-mile ISP the consumer subscribes to. This equal treatment of content is the principle of "network neutrality", which was finally codified by the FCC in 2015. The principle does not rule out paywalls erected by content providers themselves, but does prohibit last-mile ISPs from acting as gatekeepers for internet content.

In the Netflix event, *ex post* testimony makes it difficult to determine whether ISPs were acting as gatekeepers and throttling Netflix content in exchange for payment. Instead, there were numerous opportunities for both last-mile ISPs and transit ISPs (acting on Netflix's behalf) to maintain transmission capacity, but both parties failed to do so.³³ The inability to assign responsibility to last-mile ISPs even after the fact raises two questions: first, whether regulations that prohibit "gatekeeper" behaviour can be reasonably enforced, either at all or in a timely enough fashion to protect consumers; second, whether prohibiting "gatekeeper" behaviour is the key to avoiding quality of service degradation to consumers.

Rather than focus on the legal framework for interaction between last-mile ISPs, transit ISPs and content, I take the approach that renegotiation of the terms of interconnection is inevitable, and leads to unpreventable QoS degradation. I focus instead on policies that are designed to directly affect consumers—some of which, like the introduction of a publicly-owned fiber optic alternative, have already begun to be rolled out. By integrating upstream negotiations and consumer demand in a single framework, I will be able to evaluate the effectiveness of "consumer-facing" pro-competitive policies on the length of negotiations.

C Data Construction

C1. Supply Data

I draw on Netflix's throughput data, as well as MLab measurement data, CAIDA data on interconnection, and Netflix qualitative business filings data to argue that (1) quality reductions were for business reasons, not technical ones and (2) the duration of quality reductions corresponds to the duration of bargaining over interconnection.

A key document in my analysis is the public (partially redacted) version of Netflix's "Petition to Deny". The Petition to Deny is a legal document filed by Netflix to the Federal Communications Commission to argue against the application for merger of Comcast and Time Warner that was announced in mid 2014. In making its case for why these entities should not merge, Net-

³³From consultations with an economist who was with the FCC at the time the slowdown was occurring.

flix detailed the difficulties it had during the installation of its Open Connect servers in 2013. Perhaps because its goal in the Petition to Deny was to argue the dangers of too much market power, only the four largest ISPs are explicitly named in the document. While I take statements in the document about agreement timings at face value, statements that the ISPs are to blame for the slowdown will need to be evaluated.

C1.1. Ruling Out ISP Technical Difficulties

From the Netflix data it is clear that some ISPs experience a slowdown and others do not. However, it is possible that affected ISPs were suffering from general network problems and not Netflix-specific problems. I falsify this hypothesis with data from MLab, an independent research group that measures the rate of transmission of data between last-mile ISPs and transit ISPs. Since Netflix primarily uses the service of the transit ISP Cogent during this time, if slowdowns occur between Cogent and the affected ISPs but not between other transit providers and the affected ISPs, then it will be evidence of a Netflix-ISP specific problem. Figure 13 illustrates that starting in July 2013, there was a sharp drop in Netflix streaming quality to Comcast, Verizon and TimeWarner. Throughput of other content to these ISPs was not affected, and throughput of Netflix to ISPs such as Cox—a large provider with roughly 5.5% share of subscribers nationwide—was also not affected. Although the MLab data is not comprehensive in its coverage of ISPs, it strongly suggests that slowdowns occurred pairwise. Combined with the qualitative data cited in Section 2.2., it is clear that the slowdowns were business driven.

C1.2. Assigning Responsibility for the Slowdown

Although Netflix’s standing offer to install Open Connect servers was not taken up by the largest ISPs until 2013, transmission of Netflix content to the end users remained reliable until mid-2013. At that point, Netflix, the last-mile ISPs, or both Netflix and last-mile ISPs either actively precipitated a collapse in reliable transmission, or failed to proceed with status quo network upgrades.³⁴ The end result was that quality of service plummeted for the largest ISPs.

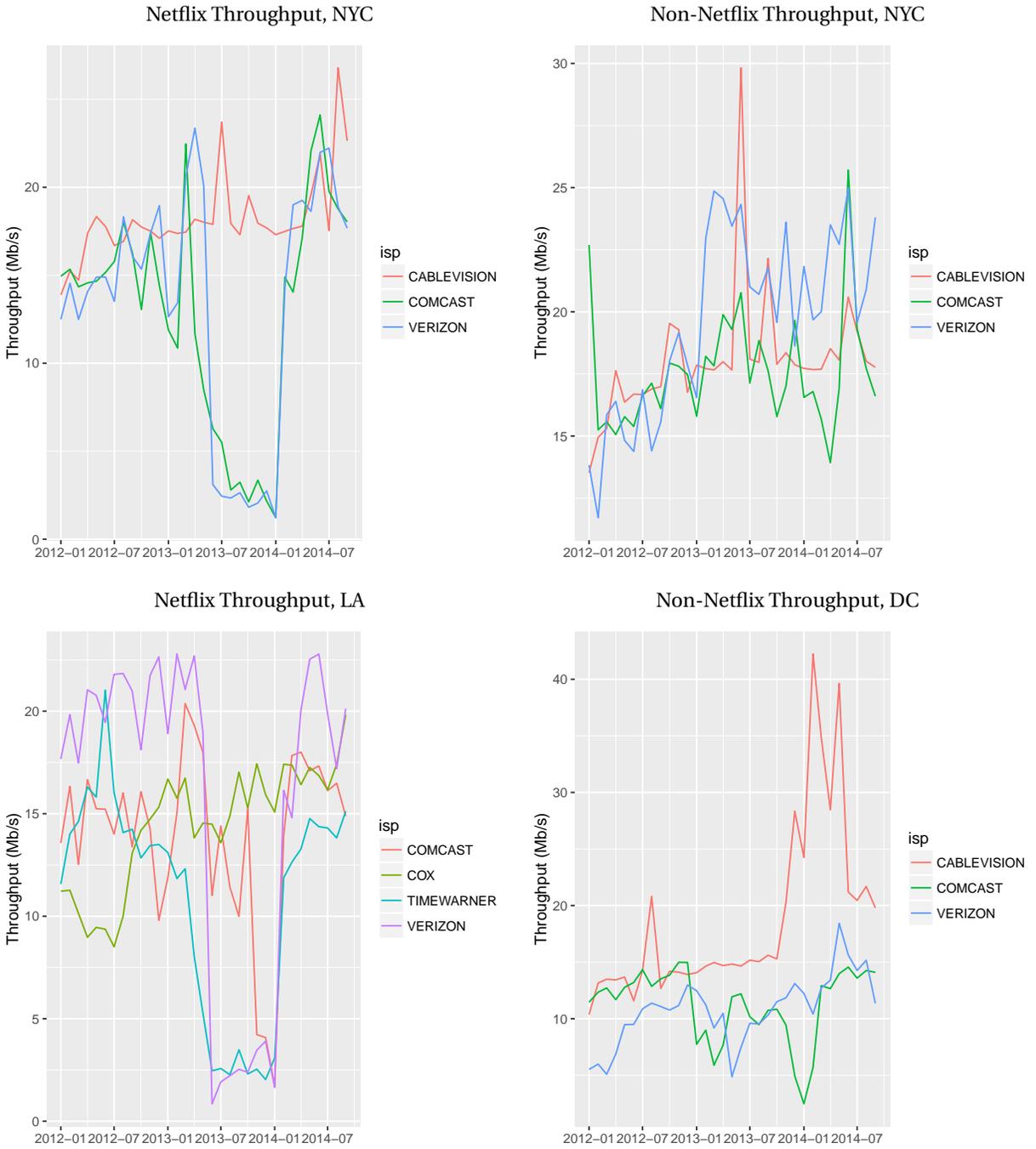
From Netflix’s point of view, it is the ISPs that were at fault: Netflix argues that Comcast, Verizon, Time Warner and AT&T “presumably made the business decision that the present discounted value of benefits from degrading the quality of the Netflix video stream to [their] subscribers was greater than the present discounted value of the costs.”³⁵

However, from CAIDA data, Netflix cancelled service with a crucial third party in late 2013. Limelight was a large CDN that Netflix had relied on to smooth delivery of its services, but after November 2013—the low point of quality degradation from Figure 1—they no longer intercon-

³⁴Harvard Business School case N9-616-007.

³⁵Petition to deny, pg. 52, paragraph 2.

Figure 13: ISP-Netflix Pairwise Throughput in NYC, LA and DC



nected. Meanwhile, Comcast claimed that Netflix’s throughput problems could be solved if they purchased more bandwidth from transit providers—a statement reported and dismissed by Netflix.³⁶

In the model, I show that Netflix had the incentive to degrade quality of service to induce faster agreement times from ISPs. I also show that some ISPs had positive *ex ante* expected payoffs from the dispute even with quality degradation. However, since the marginal consumer appears to be more elastic with respect to switching ISPs than canceling Netflix, there is no incentive for ISPs to slow down traffic.

C1.3. Constructing Durations

I assume that bargaining begins for all ISPs in 2013Q3, corresponding to the sharp drop off in quality in Figure 13.

The duration of disagreement is constructed in the following steps:

1. Using the Netflix data in Figure 1, for each ISP for which there is data, I regress throughput on a linear time trend and code a “slowdown” dummy as 1 if throughput falls below 80% of its predicted value. This is a necessary condition for ISPs to be considered as having a lengthy slowdown, and provides candidate disagreement durations.
2. If an ISP is explicitly mentioned in Netflix’s Q-10 or Petition to Deny Filings as having reached agreement or not by a certain time, I adjust the disagreement durations to reflect this (Comcast, TimeWarner, and Verizon reach agreement in 2014Q1, while AT&T does not. No other ISPs are mentioned.)³⁷
3. If any remaining ISP appears in the CAIDA data as interconnecting with OpenConnect at a certain time, I adjust the disagreement data to reflect this timing
4. All remaining ISPs are those that did not experience a quality degradation. Netflix pursued a policy of installing Open Connect infrastructure in the networks of even small ISPs, so I assume that these ISPs reached agreement with Netflix immediately.³⁸ CAIDA data indicates that some small ISPs reached agreement earlier—for instance, RCN interconnects

³⁶Statement of Ken Florance, Vice President of Content Delivery at Netflix since 2012.

³⁷In its first quarter letter to investors issued on April 21, 2014, pg. 5 paragraph 3, Netflix notes “now nearly all cable Internet households receive great quality Internet video”, implying that Time Warner Cable also concluded negotiations by the first quarter of 2014. From the same document, Netflix mentions the extremely poor streaming quality that AT&T U-Verse customers receive, and argues that “[it] is free and easy for AT&T to interconnect directly with Netflix and quickly improve their customers’ experience, should AT&T so desire.”, implying that the slowdown could be alleviated as soon as Netflix and AT&T could settle on a price. From this report it is clear that AT&T actually took longer to resolve negotiations, so that the slowdown truly indicates negotiation time and not just greater technological difficulties implementing interconnection with Open Connect.

³⁸“if an ISP has an individual market area serving a population of at least 100,000 subscribers, Netflix will install Open Connect appliances at that location at no charge to the ISP”, pg.49, paragraph 2, Netflix Petition to Deny.

with Open Connect in late 2012. I assume that ISPs that agree immediately do so in the first period of bargaining in 2013Q3.

The full set of derived durations is displayed in Table 9. I include Hughes and Wildblue, the two satellite operators in the sample, but I do not assume that these providers bargain with Netflix over interconnection since their infrastructure is not amenable to Open Connect. There is no slowdown vis-a-vis these providers however, so their bargaining duration is set to zero and they are not used in the supply estimation.

Table 9: Complete Durations Data

isp	duration	isp.1	duration.1
AT&T	4	FRONTIER	3
BRIGHTHOUSE	4	HAWAIIAN TELCOM	0
CABLE ONE	0	HUGHES	0
CABLEVISION	0	MEDIACOM	4
CENTURYLINK	3	RCN	0
CHARTER	2	SUDDENLINK	0
CINCINNATI BELL	0	TDS TELECOM	0
COMCAST	2	TIMEWARNER	2
CONSOLIDATED	0	VERIZON	2
COX	0	WIDEOPENWEST	0
EARTHLINK	0	WINDSTREAM	0
FAIRPOINT	0	FRONTIER	3